Abstract. This paper deals with asset price bubbles modeled by strict local martingales. To any strict local martingale one can associate a new measure, which is studied in detail in the first part of the paper. In the second part we determine the "default term" apparent in risk-neutral option prices if the underlying stock exhibits a bubble modeled by a strict local martingale. Results for certain path dependent options and last passage time formulas are given.

Introduction

The goal of this paper is to determine the influence of asset price bubbles on the pricing of derivatives. Asset price bubbles have been studied extensively in the economic literature looking for explanations of why they arise, but have only recently gained attention in mathematical finance by Cox and Hobson (2005), Jarrow et al. (2007, 2009, 2010), Protter and Pal (2007), and Heston et al. (2007). When an asset price bubble exists, the market price of the asset is higher than its fundamental value. From a mathematical point of view this is the case, when the stock price process is modeled by a positive strict local martingale under the equivalent local martingale measure. Here by a strict local martingale we understand a local martingale, which is not a true martingale. Strict local martingales were first studied in the context of financial mathematics by Delbaen and Schachermayer (1995). Afterwards Elworthy et al. (1997, 1999) studied some of their properties including their tail behaviour. More recently, the interest in them grew again (cf. e.g. Mijatovic and Urusov (2010)) because of their importance in the modelling of financial bubbles.

Obviously, there are options for which it does not matter whether the underlying is a strict local martingale or not, but for which well-known results still hold true without modification under the condition of no free lunch with vanishing risk (NFLVR). One example is the put option with strike \( K \geq 0 \). If the underlying is modeled by a continuous local martingale \( X \) with \( X_0 = 1 \), it is shown by Madan et al. (2008) that the risk-neutral value of the put option can be expressed in terms of the last passage time \( \rho^K_X = \sup \{ t \geq 0 \mid X_t = K \} \) of the local martingale \( X \) at level \( K \) via

\[
\mathbb{E}(K - X_T)^+ = \mathbb{E} \left( (K - X_\infty)^+ I_{\{\rho^K_X \leq T\}} \right).
\]

This formula does not require \( X \) to be a true martingale, but is also valid for strict local martingales. However this changes if we go from puts to calls. The general idea is to reduce the call case to the put case by a change of measure with Radon-Nikodym density process given by \( (X_t)_{t \geq 0} \) as
done in Madan et al. (2008) in the case where $X$ is a true martingale. However, if $X$ is a strict local martingale, this does not define a measure any more. Instead, we first have to localize the strict local martingale and can thus only define measures on stopped sub-$\sigma$-algebras. Under certain conditions on the probability space, we can then extend the so-defined consistent family of measures to a measure defined on some larger $\sigma$-field. Under the new measure the inverse of $X$ turns into a true martingale. The conditions we impose are taken from Föllmer (1972), who requires the filtration to be a standard system (cf. Definition [1.7]). This way we get an extension of Theorem 4 in Delbaen and Schachermayer (1995) to general probability spaces and càdlàg local martingales. We study the behavior of $X$ and other local martingales under the new measure.

Using these technical results we obtain decomposition formulas for some classes of European path-dependent options under the NFLVR condition. These formulas are extensions of Proposition 2 in Pal and Protter (2007), which deals with non-path-dependent options. We decompose the option value into a difference of two positive terms, of which the second one shows the influence of the stock price bubble.

Furthermore, we express the risk-neutral price of an exchange option in the presence of asset price bubbles as an expectation involving the last passage time at the strike level under the new measure. This result is similar to the formula for call options derived by Madan, Roynette and Yor (2008) or Yen and Yor (2009) for the case of inverse Bessel processes. We can further generalize their formula to the case where the candidate density process for the risk-neutral measure is only a strict local martingale. Then the NFLVR condition is not fulfilled and risk-neutral valuation fails, so that we have to work under the real-world measure. Since in this case the price of a zero coupon bond is decreasing in maturity even with an interest rate of zero, some people refer to this as a bond price bubble as opposed to the stock price bubbles discussed above. In this general setup we obtain expressions for the option value of European and American call options in terms of the last passage time and the explosion time of the deflated price process, which make some anomalies of the prices of call options in the presence of bubbles evident: European calls are not increasing in maturity any longer and the American call option premium is not equal to zero any more.

This paper is organized as follows: In the next section we study strictly positive (strict) local martingales in more detail. On the one hand, we demonstrate ways of how one can obtain strict local martingales, while on the other hand we construct the above mentioned measure associated to a càdlàg strictly positive local martingale on a general filtered probability space with a standard system as filtration. We give some examples of this construction in Section 2. In Section 3 we then apply our results to the study of asset price bubbles. After formally defining the financial market model we obtain decomposition formulas for certain classes of European path-dependent options, which show the influence of stock price bubbles on the value of the options under the NFLVR condition. In Section 4 we further study the relationship between the original and the new measure constructed in Section 1.2, which we apply in Section 5 to obtain last passage time formulas for the European and American exchange option in the presence of asset price bubbles. Moreover,
we show how this result can be applied to the real-world pricing of European and American call options. The last section contains some results about multivariate strict local martingales.

1. Càdlàg Strictly Positive Strict Local Martingales

When dealing with continuous strictly positive strict local martingales a very useful tool is the result from [6], see also Proposition 6 in [28], which states that every such process defined as the coordinate process on the canonical space of trajectories can be obtained as the inverse of a ”Doob h-transform”\(^1\) of a continuous non-negative true martingale. Conversely, any such transformation of a continuous non-negative martingale, which hits zero with positive probability, yields a strict local martingale.

The goal of this section is to extend these results to càdlàg processes and general probability spaces satisfying some extra conditions, which were introduced by Parthasarathy in [29] and used in a similar context in [13]. While the construction of strict local martingales from true martingales follows from an application of the Lenglart-Girsanov theorem, the converse theorem relies as in [6] on the construction of the Föllmer exit measure of a strictly positive local martingale as done in [13] and [25].

1.1. How to obtain strictly positive strict local martingales. Examples of continuous strict local martingales have been known for a long time, the canonical example being the inverse of a Bessel process of dimension 3. This example can be generalized to a broader class of transient diffusions, which taken in natural scale turn out to be strict local martingales, cf. [10]. More recently, further results concerning the martingale property of certain local martingales have been obtained, cf. for example [2, 26]. A natural way to construct strictly positive continuous strict local martingales is given in Theorem 1 of [6]. There, it is shown that every uniformly integrable non-negative martingale with positive probability to hit zero gives rise to a change of measure such that its inverse is a strict local martingale under the new measure. For the non-continuous case and for not necessarily uniformly integrable martingales we now give a simple extension of the just mentioned theorem from [6]:

**Theorem 1.1.** Let \((\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, Q)\) be the natural augmentation of some probability space with \(\mathcal{F} = \bigvee_{t \geq 0} \mathcal{F}_t\), i.e. the filtration \((\mathcal{F}_t)_{t \geq 0}\) is right-continuous and \(\mathcal{F}_0\) contains all \(\mathcal{F}_t\)-negligible sets for all \(t \geq 0\). Let \(X\) be a non-negative \(Q\)-martingale with \(Q(X_0 = 1) = 1\). Set \(\tau = \inf\{t \geq 0 : X_t = 0\}\) and assume that \(Q(\tau < \infty) > 0\). Furthermore, suppose that \(X\) does not jump to zero \(Q\)-almost surely. For all \(t \geq 0\), define a probability measure \(P_t\) on \(\mathcal{F}_t\) via \(P_t = X_t Q|_{\mathcal{F}_t}\); in particular, \(P_t \ll Q|_{\mathcal{F}_t}\). Assume that either \(X\) is uniformly integrable under \(Q\) or that the non-augmented probability space satisfies condition \((P)\). Then, we can extend the consistent family \((P_t)_{t \geq 0}\) to a

\(^1\)Note that we abuse the word “Doob h-transform” in this context slightly, since Doob h-transforms are normally only defined in the theory of Markov processes, cf. Appendix B.

\(^2\)Condition \((P)\) first appeared in [29] and was later used in [27]. We recall its definition in Appendix A.
measure $\mathbb{P}$ on the augmented space $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0})$. Under the measure $\mathbb{P}$ the process $X$ does never reach zero and its reciprocal $1/X$ is a strict local $\mathbb{P}$-martingale.

**Proof.** Since the underlying probability space satisfies the natural assumptions, we may choose a càdlàg version of $X$, cf. Propositions 3.1 and 3.3 in [27]. If $X$ is a uniformly integrable martingale, the measure $\mathbb{P}$ is defined on $\mathcal{F}$ by $d\mathbb{P} = X_\infty d\mathbb{Q}$. In the other case, when the probability space fulfills condition $(P)$, the existence of the measure $\mathbb{P}$ follows from Corollary 4.9 in [27]. Moreover note that

$$\mathbb{P}(\tau < \infty) = \lim_{t \to \infty} \mathbb{P}(\tau \leq t) = \lim_{t \to \infty} \mathbb{E}^{\mathbb{Q}}(\mathbb{1}_{\{\tau \leq t\}}X_t) = 0,$$

therefore the process $1/X$ is a $\mathbb{P}$-almost surely well defined semimartingale. The result now follows from Corollary 3.10 of [?] applied to $M'_t = \frac{1}{X_t}\mathbb{1}_{\{\tau > t\}}$, once we can show that $(M'_{t \wedge \tau_n}X_{t \wedge \tau_n})$ with $\tau_n = \inf\{t \geq 0 : X_t \leq \frac{1}{n}\}$ is a local $\mathbb{P}$-martingale for every $n \in \mathbb{N}$. But,

$$M'_{t \wedge \tau_n}X_{t \wedge \tau_n} = \mathbb{1}_{\{\tau > t \wedge \tau_n\}} = 1,$$

which trivially proves the martingale property. Finally, the strictness of the local martingale $1/X$ under $\mathbb{P}$ follows from

$$\mathbb{E}^{\mathbb{P}}\left(\frac{1}{X_t}\right) = Q(\tau > t) < 1$$

for $t$ large enough, since by assumption $Q(\tau < \infty) > 0$. □

Starting with a Brownian motion stopped at zero under $\mathbb{Q}$, it is easy to show that the associated strict local martingale under $\mathbb{P}$ is the reciprocal of the three-dimensional Bessel process, which is the canonical example of a strict local martingale (cf. Example 1 in [28]). Without stating the general result, the above construction is also applied in [4] to construct examples of strict local martingales with jumps related to Dunkl Markov processes on the one hand (cf. Proposition 3 in [4]) and semi-stable Markov processes on the other hand (cf. Proposition 5 in [4]). Apart from the previous, there do not seem to be any well-known examples of strict local martingales with jumps. Note, however, that one can construct an example by taking any continuous strict local martingale and multiplying it with the stochastic exponential of an independent compound Poisson process or any other independent and strictly positive jump martingale.

In the following example we construct a “non-trivial” positive strict local martingale with jumps by a shrinkage of filtration.

**Example 1.2.** Consider the well-known inverse three-dimensional Bessel process $X$ as a function of a three-dimensional standard Brownian motion $B = (B^1, B^2, B^3)$ starting from $B_0 = (1, 0, 0)$, i.e.

$$X = \frac{1}{\sqrt{(B^1)^2 + (B^2)^2 + (B^3)^2}}.$$
We define the filtrations \((\mathcal{F}_t)_{t \geq 0}\) and \((\mathcal{G}_t)_{t \geq 0}\) by \(\mathcal{F}_t = \sigma(B^1_s, B^2_s, B^3_s; s \leq t)\) and \(\mathcal{G}_t = \sigma(B^1_s, B^2_s; s \leq t)\), as well as the filtration \((\mathcal{H}_t)_{t \geq 0}\) via
\[
\mathcal{H}_t = \mathcal{F}_{\lfloor nt \rfloor / n} \lor \mathcal{G}_t = \sigma \left( B^1_s, B^2_s, s \leq t; B^3_u, u \leq \frac{|nt|}{n} \right)
\]
for some \(n \in \mathbb{N}\). It is shown in Theorem 15 of \[14\] that not only \(X\) itself is an \((\mathcal{F}_t)_{t \geq 0}\)-strict local martingale, but that also the optional projection of \(\mathcal{G}_t\) for some \(n\) for some \(\mathcal{H}_t\) is a continuous local \((\tilde{\mathcal{H}}_t)_{t \geq 0}\)-martingale. Since \(\mathcal{G}_t \subset \mathcal{H}_t \subset \mathcal{F}_t\) for \(t \geq 0\), it follows by Corollary 2 of \[14\] that then the optional projection of \(X\) onto \((\mathcal{H}_t)_{t \geq 0}\), denoted by \(\diamond X\), is also a local martingale. However, since its expectation process is decreasing, \(\diamond X\) must be a strict local martingale that jumps at \(t \in \mathbb{N}\). In fact, \(\diamond X\) is given by the explicit formula \(\diamond X_t = u(B^1_t, B^2_t, B^3_{\lfloor nt \rfloor / n}, t)\), where
\[
u(x, y, a, t) = \int_{\mathbb{R}} \left( x^2 + y^2 + z^2 \right)^{-1/2} \cdot \frac{1}{\sqrt{2\pi \left( t - \lfloor nt \rfloor / n \right)}} \exp \left( -\frac{1}{2 \left( t - \lfloor nt \rfloor / n \right)}(z - a)^2 \right) \, dz.
\]

**Remark 1.3.** As a further example, any nonnegative non-uniformly integrable \((\mathcal{F}_t)_{t \geq 0}\)-martingale \(Y\) with \(Y_0 = 1\) allows to construct a strictly positive strict local martingale \(X\) relative to a new filtration \((\tilde{\mathcal{F}}_t)_{t \geq 0}\) through a deterministic change of time: simply set
\[
X_t = \begin{cases} \frac{1}{2} \left( 1 + \frac{Y_{\lfloor nt \rfloor}}{nt} \right) & : 0 \leq t < 1 \\ \frac{1}{2} \left( 1 + \lim_{t \to \infty} Y_t \right) & : 1 \leq t \end{cases}
\]
and define \(\tilde{\mathcal{F}}_t = \mathcal{F}_t \uparrow t\) for \(t < 1\) and \(\tilde{\mathcal{F}}_t = \mathcal{F}_\infty\) for \(t \geq 1\). Since \(Y\) is not uniformly integrable, we have \(\mathbb{E}X_1 < X_0 = Y_0 = 1\) almost surely. Instead of setting \(X\) constant for \(t \geq 1\) one can also define \(X\) to behave like any other strictly positive local martingale starting from \(X_1 := \lim_{t \to \infty} Y_t\) on \([1, \infty)\). Note however that \(X\) is a true martingale on the interval \([0, 1)\).

### 1.2. From strictly positive strict local martingales to true martingales.

In the following let \((\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})\) be a filtered probability space. Furthermore, we denote by \((\mathcal{F}_t)_{t \geq 0}\) the right-continuous augmentation of \((\tilde{\mathcal{F}}_t)_{t \geq 0}\), i.e. \(\mathcal{F}_t := \tilde{\mathcal{F}}_{t+} = \bigcap_{s \geq t} \tilde{\mathcal{F}}_s\) for all \(t \geq 0\). Note, however, that the filtration is not completed with the negligible sets of \(\mathcal{F}\).

**Definition 1.4.** (cf. \[13\]) Let \((\mathcal{F}_t)_{t \in T}\) be a filtration on \(\Omega\), where \(T\) is a partially ordered non-void index set, and let \((\mathcal{F}_t)_{t \in T}\) be its right-continuous augmentation. Then \((\tilde{\mathcal{F}}_t)_{t \in T}\) is called a standard system if

- each measurable space \((\Omega, \tilde{\mathcal{F}}_t)\) is a standard Borel space, i.e. \(\tilde{\mathcal{F}}_t\) is \(\sigma\)-isomorphic to the \(\sigma\)-field of Borel sets on some complete separable metric space.
- for any increasing sequence \((t_i)_{i \in \mathbb{N}} \subset T\) and for any \(A_1 \supset A_2 \supset \cdots \supset A_i \supset \cdots\), where \(A_i\) is an atom of \(\mathcal{F}_{t_i}\), we have \(\bigcap_i A_i \neq \emptyset\).

As noted in \[27\] the filtration \(\tilde{\mathcal{F}}_t = \sigma(X_s, s \leq t)\), where \(X_t(\omega) = \omega(t)\) is the coordinate process on the space \(C(\mathbb{R}_+, \mathbb{R}_+)\) of non-explosive non-negative continuous functions, is not a standard
system. However, it will be seen below that when dealing with strict local martingales it is natural to work on the space of all non-negative continuous processes that stay at infinity once they reach infinity, \( C_\infty(\mathbb{R}_+, \mathbb{R}_+) \), where \( \mathbb{R}_+ = \mathbb{R}_+ \cup \{ \infty \} \) is endowed with the usual one-point compactification topology. As noted in example (6.3) in [13] the filtration generated by the coordinate process on this space is indeed a standard system. More generally, we have the following lemma. Recall that for any \((\mathcal{F}_t)_{t \geq 0}\)-stopping time \( \tau \) the sigma-algebra \( \mathcal{F}_{\tau^-} \) is defined as

\[
\mathcal{F}_{\tau^-} = \sigma(\mathcal{F}_0, \{ \{ \tau > t \} \cap \Gamma : \Gamma \in \mathcal{F}_t, t \geq 0 \}).
\]

**Lemma 1.5.** Let \( \Omega = D_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \) be the space of functions from \( \mathbb{R}_+ \) into \( \mathbb{R}_+^n = (\mathbb{R}_+ \cup \{ \infty \})^n, n \in \mathbb{N} \), with componentwise càdlàg paths \((\omega_i(t))_{t \geq 0}, i = 1, \ldots, n, \) that remain constant after \( \tau^- \) at the value \( \omega_i(\tau^-) \), where \( \tau = \lim_{k \to \infty} \tau_k \) and \( \tau_k = \inf\{ t \geq 0 | \exists i = 1, \ldots, n : \omega_i(t) > k \} \). We denote by \((X_t)_{t \geq 0}\) the coordinate process, i.e. \( X_t(\omega_1, \ldots, \omega_n) = (\omega_1(t), \ldots, \omega_n(t)) \), and by \((\tilde{\mathcal{F}}_t)_{t \geq 0}\) the canonical filtration generated by the coordinate process, i.e. \( \tilde{\mathcal{F}}_t = \sigma(X_s; s \leq t) \). Furthermore, set \( \mathcal{F} = \bigvee_{t \geq 0} \tilde{\mathcal{F}}_t \). Then, \((\tilde{\mathcal{F}}_t)_{t \geq 0}\) is a standard system on the space \((\Omega, \mathcal{F}, (\tilde{\mathcal{F}}_t)_{t \geq 0})\). The same is true, if we replace \( D_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \) by its subspace \( C_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \) of component-wise continuous functions allowing for explosions. Moreover, in both cases we have \( \mathcal{F}_{\tau^-} = \mathcal{F}_\tau = \mathcal{F} \).

**Proof.** We prove the claim for \( \Omega = D_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \). The case \( \Omega = C_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \) is done in a similar way. As in [9] we define a bijective mapping \( i \) from \( \mathbb{R}_+ \) to some subspace \( \mathbb{C}_+ \subset (\mathbb{R}_+^n)^\mathbb{Q} \), (where here \( \mathbb{Q} \) denotes the set of all rational numbers), via \( \omega \mapsto (X_{\tau}(\omega))_{\tau \in \mathbb{Q}} \). It is clear that \( i \) is bijective and we have \( \mathcal{F} = i^{-1}(\mathcal{B}(\mathbb{C})) \). Furthermore, a sequence \( A_1 \supset A_2 \supset \cdots \supset A_t \supset \cdots \) of atoms of \( \mathcal{F}_t = \bigcap_{s \geq t} \sigma(X_s; s \leq r) \) defines a component-wise càdlàg function on every interval \([0, \lim_{s \to \infty} t_i]\cap[0, \infty)\) for every sequence \((t_i)_{i \in \mathbb{N}} \subset \mathbb{R}_+ \), because we allow for explosions. This function can easily be extended to an element of \( D_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \). The equality \( \mathcal{F}_{\tau^-} = \mathcal{F}_\tau = \mathcal{F} \) is obvious from the definition of the respective spaces. \( \square \)

**Remark 1.6.** In the same way it can be shown that the spaces \( \Omega = D_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \) and \( C_\infty(\mathbb{R}_+, \mathbb{R}_+^n) \) of componentwise càdlàg resp. continuous functions from \( \mathbb{R}_+ \) into \( \mathbb{R}_+^n = (\mathbb{R} \cup \{ \infty \} \cup \{ -\infty \})^n \), that remain constant after \( \tau^- \), where \( \tau = \lim_{k \to \infty} \tau_k \) and \( \tau_k = \inf\{ t \geq 0 | \exists i = 1, \ldots, n : |\omega_i(t)| > k \} \), with their canonical filtrations generated by the coordinate process are also standard spaces.

**Lemma 1.7.** (cf. [13], Remark 6.1) Let \((\tilde{\mathcal{F}}_t)_{t \geq 0}\) be a standard system on \( \Omega \). Then for any increasing sequence \((\tau_n)_{n \in \mathbb{N}}\) of \((\mathcal{F}_t)\)-stopping times the family \((\mathcal{F}_{\tau_n^-})_{n \in \mathbb{N}}\) is also a standard system.

**Notation:** When working on the subspace \((\Omega, \mathcal{F}_{\tau^-})\) of \((\Omega, \mathcal{F})\), where \( \tau \) is some \((\mathcal{F}_t)\)-stopping time, we must restrict the filtration to \((\mathcal{F}_{t \wedge \tau^-})_{t \geq 0}\), where with a slight abuse of notation we set \( \mathcal{F}_{t \wedge \tau^-} := \mathcal{F}_t \cap \mathcal{F}_{\tau^-} \). In the following we may also write \((\mathcal{F}_t)_{0 \leq t < \tau}\) for the filtration on \((\Omega, \mathcal{F}_{\tau^-}, \mathbb{P})\).

Working with standard systems will allow us to derive for every strictly positive strict local \( \mathbb{P}\)-martingale the existence of a measure \( \mathbb{Q} \) on \((\Omega, \mathcal{F}_{\tau^-}, (\mathcal{F}_t)_{0 \leq t < \tau})\), such that the reciprocal of the strict local \( \mathbb{P}\)-martingale is a true \( \mathbb{Q}\)-martingale. In Section 3 we will use this result to reduce
calculations involving strict local martingales to the much easier case of true martingales.

From Theorem 4 in [6] and Proposition 6 [28] we know that every continuous local martingale understood as the canonical process on $C(\mathbb{R}_+,\mathbb{R}_+)$ gives rise to a new measure under which its inverse turns into a true martingale. The following theorem is an extension of this result to more general probability spaces and càdlàg processes. Its proof relies on the construction of the Föllmer measure, cf. [13]; nevertheless we will give a detailed proof, since it is essential for the rest of the paper.

**Theorem 1.8.** Let $\left(\Omega,\mathcal{F},(\mathcal{F}_t)_{t\geq0},\mathbb{P}\right)$ be a filtered probability space and assume that $(\mathcal{F}_t)_{t\geq0}$ is a standard system. Let $X$ be a càdlàg local martingale on the space $(\Omega,\mathcal{F},(\mathcal{F}_t)_{t\geq0},\mathbb{P})$ with values in $(0,\infty)$ and $X_0 = 1$ $\mathbb{P}$-almost surely. We define $\tau_n^X := \inf\{t \geq 0 : X_t > n\} \land n$ and $\tau^X = \lim_{n \to \infty} \tau_n^X$.

Then there exists a unique probability measure $Q$ on $(\Omega,\mathcal{F}_{\tau^X},(\mathcal{F}_{t\land\tau^X})_{t\geq0})$, such that $\frac{1}{X}$ is a $Q$-martingale up to time $\tau^X$ and $Q|_{\mathcal{F}_t\cap\mathcal{F}_{\tau^X}} \gg \mathbb{P}|_{\mathcal{F}_t\cap\mathcal{F}_{\tau^X}}$ for all $t \geq 0$ with Radon-Nikodym derivative given by

$$\frac{dQ}{d\mathbb{P}}\bigg|_{\mathcal{F}_t \cap \mathcal{F}_{\tau^X}} = \frac{1}{X_t} \mathbb{1}_{\{t < \tau^X\}}.$$ 

**Proof.** First, note that $\tau_n^X$ is an $(\mathcal{F}_t)_{t\geq0}$-stopping time and the process $(X_{t\land\tau^X})_{t\geq0}$ is a uniformly integrable $(\mathcal{F}_t)_{t\geq0}$-martingale for all $n \in \mathbb{N}$. Furthermore, $\mathbb{P}(\tau^X = \infty) = 1$, since a positive càdlàg local martingale does not explode almost surely. We define on $(\Omega,\mathcal{F}_{\tau^X})$ the probability measure $\tilde{Q}_n$ via $\tilde{Q}_n = X_{\tau^X} \mathbb{P}|_{\mathcal{F}_{\tau^X}}$ for all $n \in \mathbb{N}$. The family $(\tilde{Q}_n)_{n\in\mathbb{N}}$ constitutes a consistent family of probability measures on $(\mathcal{F}_{\tau^X})_{n\geq1}$: If $A \in \mathcal{F}_{\tau^X}$, then

$$\tilde{Q}_{n+k}(A) = \mathbb{E}^\mathbb{P}\left(X_{\tau^X_{n+k}} \mathbb{1}_A\right) = \mathbb{E}^\mathbb{P}(X_{\tau^X_n} \mathbb{1}_A) = \tilde{Q}_n(A),$$

i.e. $\tilde{Q}_{n+k}|_{\mathcal{F}_{\tau^X_n}} = \tilde{Q}_n$ for all $n, k \in \mathbb{N}$. This induces a sequence of consistently defined measures $(\tilde{Q}_n)_{n\in\mathbb{N}}$ on the sequence $(\mathcal{F}_{\tau^X_n})_{n\in\mathbb{N}}$, which is a standard system by Lemma 1.7. Note that $\mathcal{F}_{\tau^X} = \bigvee_{n\geq1} \mathcal{F}_{\tau^X_n}$, since $(\tau_n^X)_{n\geq1}$ is increasing. We can thus apply Theorem 3.2 together with Theorem 4.1 in Chapter V of [29], cf. also Theorem 6.2 in [13], which yield the existence of a unique measure $Q$ on $(\Omega,\mathcal{F}_{\tau^X},(\mathcal{F}_{t\land\tau^X})_{t\geq0})$ such that $Q|_{\mathcal{F}_{\tau^X} \cap \mathcal{F}_{t\land\tau^X}} = \tilde{Q}_n|_{\mathcal{F}_{\tau^X_n}}$. Moreover, since $\{\tau_n^X < \tau_m^X\} \in \mathcal{F}_{\tau^X_m}$,

$$Q(\tau_n^X < \tau^X) = \lim_{m \to \infty} Q(\tau_n^X < \tau_m^X) = \lim_{m \to \infty} \tilde{Q}_m(\tau_n^X < \tau_m^X) = \lim_{m \to \infty} \mathbb{E}^\mathbb{P}\left(\mathbb{1}_{\{\tau_n^X < \tau_m^X\}} X_{\tau_m^X}\right)$$

$$= \lim_{m \to \infty} \mathbb{E}^\mathbb{P}\left(\mathbb{1}_{\{\tau_n^X < \tau_m^X\}} X_{\tau_n^X}\right) = \mathbb{E}^\mathbb{P}\left(\mathbb{1}_{\{\tau_n^X < \tau^X\}} X_{\tau_n^X}\right) = \mathbb{E}^\mathbb{P}\left(X_{\tau_n^X}\right) = 1,$$

i.e., $1/X$ does not jump to infinity under $Q$. Therefore, if $\Lambda_n \in \mathcal{F}_{\tau^X_n}$, then

$$Q(\Lambda_n) = Q(\Lambda_n \cap \{\tau^X > \tau_n^X\}) = \lim_{m \to \infty} Q(\Lambda_n \cap \{\tau_m^X > \tau_n^X\}) = \lim_{m \to \infty} \mathbb{E}^\mathbb{P}\left(X_{\tau_m^X} \mathbb{1}_{\Lambda_n} \mathbb{1}_{\{\tau_m^X > \tau_n^X\}}\right)$$

$$= \lim_{m \to \infty} \mathbb{E}^\mathbb{P}\left(X_{\tau_n^X} \mathbb{1}_{\Lambda_n} \mathbb{1}_{\{\tau_m^X > \tau_n^X\}}\right) = \mathbb{E}^\mathbb{P}\left(X_{\tau_n^X} \mathbb{1}_{\Lambda_n}\right) = \tilde{Q}_n(\Lambda_n).$$

Therefore, $Q|_{\mathcal{F}_{\tau^X_n}} = \tilde{Q}_n$ for all $n \in \mathbb{N}$. 
Now let \( S \) be an \((\mathcal{F}_t)_{t \geq 0}\) stopping time. Note that \( \{ S < \tau^X_n \} \in \mathcal{F}_S \) and \( \{ S < \tau^X_n \} \in \mathcal{F}_{\tau^X_n} \). Thus, \( (1) \)
\[
Q(S < \tau^X_n) = \hat{Q}_n(S < \tau^X_n) = \mathbb{E}^P \left( \mathbb{1}_{\{S < \tau^X_n\}} X_{\tau^X_n} \right) = \mathbb{E}^P \left( \mathbb{1}_{\{S < \tau^X_n\}} \mathbb{E}^P(X_{\tau^X_n} | \mathcal{F}_S) \right) = \mathbb{E}^P \left( \mathbb{1}_{\{S < \tau^X_n\}} X_S \right).
\]
Since \( P(\tau^X_n < \tau^X = \infty) = 1 \), taking the limit as \( n \to \infty \) in equation \( (1) \) yields \( (2) \)
\[
Q(S < \tau^X) = \mathbb{E}^P \left( \mathbb{1}_{\{S < \tau^X\}} X_S \right).
\]

Applied to the stopping time \( S_A := S \mathbb{1}_A + \infty \mathbb{1}_{A^c} \), where \( A \in \mathcal{F}_S \), this gives
\[
Q(S < \tau^X, A) = \mathbb{E}^P \left( \mathbb{1}_{A \cap \{S < \tau^X\}} X_S \right).
\]
Especially, if \( S \) is finite \( P \)-almost surely, then \( Q(S < \tau^X, A) = \mathbb{E}^P(X_S \mathbb{1}_A) \) for \( A \in \mathcal{F}_S \).

Now assume that \( A \in \mathcal{F}_t \cap \mathcal{F}_{\tau^X_\infty} \) with \( Q(A) = 0 \). Then,
\[
0 = Q(A) \geq Q(t < \tau^X, A) = \mathbb{E}^P \left( \mathbb{1}_A X_t \right) \geq 0 \quad \text{for} \quad X_t > 0 \stackrel{P-a.s.}{\Rightarrow} \mathbb{P}(A) = 0.
\]

This shows that \( Q|_{\mathcal{F}_t \cap \mathcal{F}_{\tau^X_\infty}} \gg \mathbb{P}|_{\mathcal{F}_t \cap \mathcal{F}_{\tau^X_\infty}} \) for all \( t \geq 0 \). If \( A \in \mathcal{F}_t \cap \mathcal{F}_{\tau^X_\infty} \), then
\[
\mathbb{P}(A) = \lim_{n \to \infty} \mathbb{P}(A \cap \{ t < \tau^X_n \}) = \lim_{n \to \infty} \mathbb{E}^Q \left( \mathbb{1}_A \mathbb{1}_{\{t < \tau^X_n\}} \frac{1}{X_{\tau^X_n}} \right) = \lim_{n \to \infty} \mathbb{E}^Q \left( \mathbb{1}_A \mathbb{1}_{\{t < \tau^X\}} \frac{1}{X_t} \right).
\]

Therefore, \( \frac{d\mathbb{P}}{dQ}\big|_{\mathcal{F}_t \cap \mathcal{F}_{\tau^X_\infty}} = \frac{1}{X_t} \mathbb{1}_{\{t < \tau^X\}} \) for all \( t \geq 0 \).

Finally, note that because \( (X^X_{\tau^X_n})_{t \geq 0} \) is a positive finite-valued uniformly integrable \( \mathbb{P} \)-martingale for all \( n \in \mathbb{N} \), \( \mathbb{P}|_{\mathcal{F}_{\tau^X_n}} \sim Q|_{\mathcal{F}_{\tau^X_n}} \) and
\[
d\mathbb{P}|_{\tau^X_n} = \frac{1}{X_{\tau^X_n}} dQ|_{\tau^X_n} \quad \iff \quad \frac{dQ}{d\mathbb{P}}|_{\tau^X_n} = X_t \wedge \tau^X \quad \forall \ t \geq 0.
\]

Thus,
\[
\mathbb{E}^Q \left( \frac{1}{X_t \wedge \tau^X} \bigg| \mathcal{F}_s \right) = \mathbb{E}^P \left( \frac{1}{X_s \wedge \tau^X} \cdot \frac{X_{t \wedge \tau^X}}{X_{s \wedge \tau^X}} \bigg| \mathcal{F}_s \right) = \frac{1}{X_{s \wedge \tau^X}}
\]
for \( s \leq t \), i.e. \( \frac{1}{X} \) is a local martingale up to time \( \tau^X \).

**Corollary 1.9.** Under the assumptions of Theorem 1.8, \( X \) is a strict local \( \mathbb{P} \)-martingale, if and only if \( Q(\tau^X < \infty) > 0 \).

**Proof.** It follows directly from equation \( (2) \) that \( Q(t < \tau^X) = \mathbb{E}^P X_t \), which is smaller than 1 for some \( t \), iff \( X \) is a strict local martingale under \( \mathbb{P} \). \( \square \)

**Remark 1.10.** Corollary 1.9 makes clear why we cannot work with the natural augmentation of \( (\tilde{\mathcal{F}}_t)_{t \geq 0} \). Indeed, we have \( A_n := \{ \tau^X \leq n \} \in \mathcal{F}_n \cap \mathcal{F}_{\tau^X_\infty} \) and \( P(A_n) = 0 \) for all \( n \in \mathbb{N} \), while \( Q(A_n) > 0 \) for some \( n \) if \( X \) is a strict local \( \mathbb{P} \)-martingale. However, it is in general rather inconvenient to work without any augmentation, especially if one works with an uncountable number of stochastic processes as for example in an incomplete market situation. For this reason a new kind
of augmentation - called the \((\tau_n^X)\)-natural augmentation - is introduced in \cite{21}, which is suitable for the change of measure from \(P\) to \(Q\) undertaken here. Since for the financial applications in the second part of this paper the setup introduced above is already sufficient, we do not bother about this augmentation here and refer the interested reader to \cite{21} for more technical details.

In the following we extend the measure \(Q\) in an arbitrary way from \(\mathcal{F}_{\tau^{-}}\) to \(\mathcal{F}_{\infty} = \bigvee_{t \geq 0} \tilde{\mathcal{F}}_t\). For notational convenience we assume that \(\mathcal{F} = \mathcal{F}_{\infty}\). In fact it is always possible to extend a probability measure from \(\mathcal{F}_{\tau^{-}}\) to \(\mathcal{F}\): since \((\Omega, \tilde{\mathcal{F}}_t)\) is a standard Borel space for every \(t \geq 0\) and \((\Omega, \mathcal{F}_{\tau^{-}})\) is a standard Borel space for all \(n \in \mathbb{N}\) by Lemma 1.7, it follows from Theorem 4.1 in \cite{29} that \((\Omega, \mathcal{F})\) and \((\Omega, \mathcal{F}_{\tau^{-}})\) are also standard Borel spaces. Especially, they are countably generated which allows us to apply Theorem 3.1 of \cite{12} that guarantees an extension of \(Q\) from \(\mathcal{F}_{\tau^{-}}\) to \(\mathcal{F}\). Moreover, it does not matter for the results how we extend it, because all events that happen with positive probability under \(P\) take place before time \(\tau^X\) under \(Q\) almost surely. However, if \(Y\) is any process on \((\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, P)\), then \(Y_t\) is only defined on \(\{t < \tau^X\}\) under \(Q\). Especially, if \(Y\) is a \(P\)-semimartingale, then \(Y_n^X\) is a \(Q\)-semimartingale for each \(n \in \mathbb{N}\) as follows from Girsanov’s theorem, since \(Q|\mathcal{F}_{\tau^{-}} \sim P|\mathcal{F}_{\tau^{-}}\). Therefore, \(Q\) is a \(Q\)-semimartingale on the stochastic interval \(\bigcup_{n \in \mathbb{N}} [0, \tau_n^X]\) or a “semimartingale up to time \(\tau^X\)” in the terminology of \cite{16}. We note that in general it may not be possible to extend \(Y\) to the whole positive real line under \(Q\) in such a way that \(Y\) remains a semimartingale. Indeed, according to Proposition 5.8 of \cite{16} such an extension is possible if and only if \(Y_{\tau^{-}}\) exists in \(\mathbb{R}_+\) \(Q\)-almost surely. We define the process \(\tilde{Y}\) as

\[
\tilde{Y}_t = \begin{cases} 
Y_t & : t < \tau^X \\
\lim\inf_{s \to \tau^X, s < \tau^X, s \in Q} Y_s & : \tau^X \leq t < \infty
\end{cases}
\]

Note that \(\tilde{Y}_t = Y_t\) on \(\{t < \tau^X\}\). The above definition specifies an extension of the process \(Y\), which is a priori only defined up to time \(\tau^X\), to the whole positive real line. In the following we will work with this extension.

**Lemma 1.11.** Under the assumptions of Theorem 1.8 we have \(\frac{1}{X_t} = \frac{1}{X_t^\tau} \mathbb{I}_{\{t < \tau^X\}}\). Furthermore, the process \(\left(\frac{1}{X_t}\right)_{t \geq 0}\) is a true \(Q\)-martingale for any extension of \(Q\) from \(\mathcal{F}_{\tau^{-}}\) to \(\mathcal{F}\).

**Proof.** First note that

\[
\limsup_{n \to \infty} \frac{1}{X_{t \wedge \tau_n^X}} = \limsup_{n \to \infty} \left(\frac{1}{X_t} \mathbb{I}_{\{t < \tau_n^X\}} + \frac{1}{X_{\tau_n^X}} \mathbb{I}_{\{t \geq \tau_n^X\}}\right) \leq \frac{1}{X_t} \limsup_{n \to \infty} \frac{1}{X_t} \mathbb{I}_{\{t < \tau^X\}} + \limsup_{n \to \infty} \frac{1}{X_{\tau_n^X}} \mathbb{I}_{\{t \geq \tau_n^X\}} = \frac{1}{X_t} \mathbb{I}_{\{t < \tau^X\}}
\]

and

\[
\liminf_{n \to \infty} \frac{1}{X_{t \wedge \tau_n^X}} = \liminf_{n \to \infty} \left(\frac{1}{X_t} \mathbb{I}_{\{t < \tau_n^X\}} + \frac{1}{X_{\tau_n^X}} \mathbb{I}_{\{t \geq \tau_n^X\}}\right) \geq \liminf_{n \to \infty} \frac{1}{X_t} \mathbb{I}_{\{t < \tau^X\}} = \frac{1}{X_t} \mathbb{I}_{\{t < \tau^X\}} \text{ Q.a.s.}
\]

Thus, \(\frac{1}{X_t} = \frac{1}{X_t^\tau} \mathbb{I}_{\{t < \tau^X\}}\). Furthermore,

\[
0 \leq \frac{1}{X_{\tau^X}} \mathbb{I}_{\{\tau^X < \infty\}} = \lim_{k \to \infty} \frac{1}{X_{\tau^X}} \sum_{k} \mathbb{I}_{\{t < k\}} = \lim_{k \to \infty} \lim_{n \to \infty} \frac{1}{X_{\tau^X}} \mathbb{I}_{\{t < k\}} = 0
\]
implies that $X_{\tau^X} = \infty$ on $\{\tau^X < \infty\}$ $Q$-almost surely. From Theorem 1.8 we know that $\frac{1}{X_{t \wedge \tau^X_n}}$ is a true $Q$-martingale for all $n \in \mathbb{N}$. By the definition of $\tau^X_n$ we have for any integer $n \geq t$:

$$X_{t \wedge \tau^X_n} = \tilde{X}_{t \wedge \tau^X_n} = X_{t \wedge \inf\{s \geq 0: \tilde{X}_s > n\}} \geq \tilde{X}_t \wedge 1 \implies \frac{1}{X_{t \wedge \tau^X_n}} \leq \frac{1}{\tilde{X}_t \wedge 1} = 1 \vee \frac{1}{\tilde{X}_t}.$$

Because

$$E^Q\left(\frac{1}{\tilde{X}_t}\right) = E^Q\left(\liminf_{n \to \infty} \frac{1}{X_{t \wedge \tau^X_n}}\right) \leq \liminf_{n \to \infty} E^Q\left(\frac{1}{X_{t \wedge \tau^X_n}}\right) = 1,$$

the dominated convergence theorem implies that for all $0 \leq s \leq t$

$$E^Q\left(\frac{1}{\tilde{X}_t} \mid F_s\right) = E^Q\left(\lim_{n \to \infty} \frac{1}{X_{t \wedge \tau^X_n}} \mid F_s\right) = \lim_{n \to \infty} E^Q\left(\frac{1}{X_{t \wedge \tau^X_n}} \mid F_s\right) = \lim_{n \to \infty} \frac{1}{X_{s \wedge \tau^X_n}} = \frac{1}{X_s}.$$

□

To simplify notation we identity in the following the process $X$ with $\tilde{X}$.

2. Examples

In this section we shed new light on some known examples of strict local martingales by applying the theory from the last section for illustration.

2.1. Continuous local martingales. For the following examples we work on the pathspace $C_\infty(\mathbb{R}^+, \mathbb{R}^+)$ with $X$ denoting the coordinate process and $(F_t)_{t \geq 0}$ being the right-continuous augmentation of the canonical filtration generated by the coordinate process.

2.1.1. Scaled transient diffusions. Let $(X_t)_{0 \leq t \leq \zeta}$ be a regular transient homogeneous diffusion on $(0, \infty)$, i.e.

$$X_t = x_0 + \int_0^{t \wedge \zeta} b(X_s)ds + \int_0^{t \wedge \zeta} \sigma(X_s)dW_s,$$

where $\zeta = \inf\{t \geq 0: X_t = 0 \vee X_t = \infty\} = \tau^X \wedge \tau^{1/X}$ and as before $X_0 = 1$. Then a scale function for $X$ is

$$s(x) = \int_c^x \exp\left(-2 \int_c^u \frac{b(v)}{\sigma^2(v)} dv\right) du$$

for some $c > 0$ such that $s(1) = -1$. We assume that $s(0) = -\infty$ and $s(\infty) = 0$ and that 0 is an entrance boundary for $X$. Under these assumptions it was shown in [10] that $(-s(X_t))_{0 \leq t \leq \zeta}$ is a strict local martingale. Indeed, we have

$$d(-s(X_t)) = -s'(X_t)\sigma(X_t)dW_t.$$

The behaviour of $\left(\frac{-1}{s(X_t)}\right)_{0 \leq t \leq \zeta}$ under $Q$ is given by

$$d\left(\frac{-1}{s(X_t)}\right) = \frac{s'(X_t)\sigma(X_t)}{s^2(X_t)} dW_t^Q.$$
where $W^Q$ is a $Q$-Brownian motion. Indeed, using that $\frac{dp}{dQ}\bigg|_{\mathcal{F}_t} = \frac{1}{-s(X_t)}$, Girsanov’s theorem implies that

$$W_t^Q - \int_0^t s(X_t)d\left\langle \frac{1}{s(X)}, W^Q \right\rangle_t = W_t^Q + \int_0^t \frac{s'(X_t)\sigma(X_t)}{s(X_t)}dt = W_t$$

is actually a $P$-Brownian motion. Furthermore,

$$d(-s(X_t)) = d\left(-\frac{1}{-1/s(X_t)}\right) = -s^2(X_t)d\left(\frac{-1}{s(X_t)}\right) - s^3(X_t)d\left(\frac{-1}{s(X)}\right)_t$$

$$= -s'(X_t)\sigma(X_t)dW_t^Q - \frac{[s'(X_t)]^2\sigma^2(X_t)}{s(X_t)}dt = -s'(X_t)\sigma(X_t)dW_t.$$

2.1.2. Exponential local martingales. In this subsection we consider positive local martingales of the form

$$dX_t = X_t b(Y_t)dW_t, \quad X_0 = 1,$$

where $Y$ is assumed to be a (possibly explosive) diffusion following

$$dY_t = \mu(Y_t)dt + \sigma(Y_t)dW_t, \quad Y_0 = y \in \mathbb{R},$$

with a $P$-Brownian motion $W$. Exponential local martingales of this type are further studied in [26]. Assuming that $\sigma(x) \neq 0$ for all $x > 0$ and that

$$\frac{1}{\sigma^2}, \quad \frac{\mu}{\sigma^2}, \quad \frac{b^2}{\sigma^2} \in L^1_{loc}(\mathbb{R}_+), \quad \int_0^t b^2(Y_u)du < \infty \quad \forall \ t \geq 0 \quad P\text{-a.s.}$$

a strictly positive weak solution of the given SDEs exists. Under $Q$ the dynamics of $\frac{1}{X_t}$ are

$$d\left(\frac{1}{X_t}\right) = -\frac{b(Y_t)}{X_t}dW_t^Q$$

for a $Q$-Brownian motion $W^Q$, and the $Q$-dynamics of $Y_t$ are

$$dY_t = [\mu(Y_t) + \sigma(Y_t)b(Y_t)]dt + \sigma(Y_t)dW_t^Q.$$

Notably, the criterion whether $X$ is a strict local or a true $P$-martingale from [26], Theorem 2.1, is deterministic and only involves the functions $b, \sigma$ and $\mu$ via the scale function of the original diffusion $Y$ under $P$ and an auxiliary diffusion $\tilde{Y}$, whose dynamics are identical with the $Q$-dynamics of $Y$ stated above.

---

Footnote: In [26] the conditions on $b$ are less restrictive and the exponential local martingale $X$ may hit zero, but in our setup the local martingale $X$ needs to be strictly positive.
2.1.3. Diffusions in natural scale. We now further specify the setting from 2.1.1 to the case, where
\[ dX_t = \sigma(X_t) dW_t, \quad X_0 = 1, \]
is already a \( P \)-local martingale, assuming that \( \sigma(x) \) is locally bounded and bounded away from zero for \( x > 0 \) and \( \sigma(0) = 0 \). Using the results from [8], we know that \( X \) is strictly positive, whenever
\[ \int_0^1 \frac{x}{\sigma^2(x)} dx = \infty, \]
which we shall assume in the following. Furthermore, \( X \) is a strict local martingale, iff
\[ \int_1^\infty \frac{x}{\sigma^2(x)} dx < \infty. \]
We know that \( \frac{1}{X} \) is a \( Q \)-martingale, where \( \frac{dP}{dQ} \bigg|_{F_t} = \frac{1}{X_t} \), with decomposition
\[ d\left( \frac{1}{X_t} \right) = -\frac{\sigma(X_t)}{X_t^2} dW_t^Q = \frac{1}{X_t} dW_t^Q \]
for a \( Q \)-Brownian motion \( W^Q \) and \( \sigma(y) := -y^2 \cdot \sigma\left( \frac{1}{y} \right) \). Note that
\[ \int_1^\infty \frac{y}{\sigma^2(y)} dy = \int_0^1 \frac{x}{\sigma^2(x)} dx = \infty, \]
which confirms that \( \frac{1}{X} \) is a true \( Q \)-martingale. We see that, if \( X \) is a strict local martingale under \( P \), then
\[ \int_0^1 \frac{y}{\sigma^2(y)} dy = \int_1^\infty \frac{x}{\sigma^2(x)} dx < \infty, \]
i.e. \( \frac{1}{X} \) hits zero in finite time \( Q \)-almost surely.

2.2. Jump example. Let \( \Omega = D(\mathbb{R}_+, \mathbb{R}) \) with \( (\xi_t)_{t \geq 0} \) denoting the coordinate process and \( (\mathcal{F}_t)_{t \geq 0} \) being the right-continuous augmentation of the canonical filtration generated by the coordinate process. Assume that under \( P \), \( (\xi_t)_{t \geq 0} \) is a one-dimensional Lévy process with \( \xi_0 = 0 \), \( \mathbb{E}^P \exp(b\xi_t) = \exp(t\rho(b)) < \infty \) for all \( t \geq 0 \) and characteristic exponent
\[ \Psi(\lambda) = ia\lambda + \frac{1}{2} \sigma^2 x^2 + \int_{\mathbb{R}} \left( 1 - e^{i\lambda x} + i\lambda x I_{\{|x|<1\}} \right) \pi(dx), \]
where \( a \in \mathbb{R}, \sigma^2 \geq 0 \) and \( \pi \) is a positive measure on \( \mathbb{R} \setminus \{0\} \) such that \( \int (1 \wedge |x|^2) \pi(dx) < \infty \). Define
\[ X_t = Y_t^b \exp\left( -\rho(b) \int_0^t \frac{ds}{Y_s} \right), \]
where \( (Y_t)_{t \geq 0} \) is a semi-stable Markov process, i.e. \( \left( \frac{1}{c} Y_{ct}^{(x)} \right)_{t \geq 0} \overset{(d)}{=} \left( (Y_{ct}^{(x^{-1})})_{t \geq 0} \right) \) for all \( c > 0 \), implicitly defined via
\[ \exp(\xi_t) = Y_{t_0}^c \exp(\xi_u) ds. \]

\[ \text{This example is taken from [3]. However, we corrected a small mistake concerning the time-scaling.} \]
Following [4], \((X_t)_{t \geq 0}\) is a strict local martingale if \(b\) satisfies
\[-a + \int_{|x| > 1} x \pi(dx) \geq 0, \quad -a + b \sigma^2 - \int_{|x| < 1} x (1 - e^{bx}) \pi(dx) + \int_{|x| > 1} xe^{bx} \pi(dx) < 0.\]
Furthermore, under the new measure \(Q\) the process
\[\frac{1}{X_t} = Y_t^{-b} \exp \left( \rho(b) \int_0^t ds \frac{Y_s}{Y_0} \right)\]
is a true martingale, where now \((\xi_t)_{t \geq 0}\) has characteristic component \(\tilde{\Psi}\) with
\[\tilde{\Psi}(u) = \Psi(u - ib) - \Psi(-ib),\]

3. Application to Financial Bubbles I: Decomposition Formulas

In this section we apply our results to option pricing in the presence of strict local martingales. For this, let \(X\) be a strictly positive càdlàg local martingale with \(X_0 = 1\) on a probability space \((\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, P)\) that satisfies the assumptions of Theorem 1.8, i.e. the filtration is a standard system. The stopping times \(\tau_{X}^{n}, n \in \mathbb{N}\), and \(\tau_{X}\) are defined as in sections 1.2 and 4. Furthermore, we denote by \(Q\) any extension of the measure \(Q\) constructed in Theorem 1.8 to \((\Omega, \mathcal{F})\).

We consider a financial market model which satisfies the NFLVR property as defined in [7]. We denote by \(P\) the equivalent local martingale measure (ELMM). Assuming that the interest rate equals zero, we interpret \(X\) as the (discounted) stock price process, which is a local martingale under \(P\). In this context, the question of whether \(X\) is a strict local or a true \(P\)-martingale determines whether there exists a stock price bubble. If \(X\) is a strict local \(P\)-martingale, the fundamental value of the asset (given by the conditional expectation) deviates from its actual market price \(X\). Several authors like Cox and Hobson [5], Jarrow et al. [18, 19] or Pal and Protter [28] have interpreted this as the existence of a stock price bubble, which we formally define as follows:

**Definition 3.1.** With the previous notation the *asset price bubble* for the stock price process \(X\) between time \(t \geq 0\) and time \(T \geq t\) is equal to the \(\mathcal{F}_T\)-measurable random variable
\[\gamma_X(t, T) := X_t - \mathbb{E}^P(X_T|\mathcal{F}_t).\]

**Remark 3.2.** For \(t = 0\) we recover the default function \(\gamma_X(0, T) = X_0 - \mathbb{E}^P X_T\) of the local martingale \(X\), which was introduced in [11]. In [10, 11] the authors derive several expressions for the default function in terms of the first hitting time, the local time and the last passage time of the local martingale.

In [28, Proposition 7], the price of a non-path-dependent option written on a stock, whose price process is a (strict) local martingale, is decomposed into a "normal" ("non-bubble") term and a default term. In the following we give an extension of this theorem to a certain class of path-dependent options. For this let us introduce the following notation for all \(k \in \mathbb{N}\):
\[
\mathbb{R}_+^k = \{x \in \mathbb{R}^k : x_l \geq 0, \ l = 1, \ldots, k\}, \quad \mathbb{R}_{++}^k = \{x \in \mathbb{R}^k : x_l > 0, \ l = 1, \ldots, k\}.
\]
Theorem 3.3. Let $0 \leq t_1 < t_2 < \cdots < t_n \in \mathbb{R}_+$ and consider any Borel-measurable non-negative function $h : \mathbb{R}_+^n \to \mathbb{R}_+$. Set $g(x) := x_n \cdot h \left( \frac{1}{x_1}, \ldots, \frac{1}{x_n} \right)$ for all $x = (x_1, \ldots, x_n) \in \mathbb{R}_+^n$. Then

$$
\mathbb{E}^P h(X_{t_1}, \ldots, X_{t_n}) = \mathbb{E}^Q \left( g \left( \frac{1}{X_{t_1}}, \ldots, \frac{1}{X_{t_n}} \right) \mathbb{I}_{\{\tau^X > t_n\}} \right).
$$

Now suppose that the following limits exist in $\mathbb{R}_+$ for $y_i \in \mathbb{R}_{++}$, $i = 1, \ldots, n - 1$:

$$
\lim_{|z| \to 0} g(y_1, \ldots, y_k; z_1, \ldots, z_{n-k}) =: \eta_k(y_1, \ldots, y_k), \quad k = 1, \ldots, n - 1,
$$

$$
\lim_{|z| \to 0} g(z_1, \ldots, z_n) =: \eta_0.
$$

Define $\tilde{g} : A \to \mathbb{R}_+$ as the extension of $g$ from $\mathbb{R}_+^n$ to $A \subset \mathbb{R}_+$, where $A$ is defined as

$$
A := \{ x \in \mathbb{R}_+^n : \text{if } x_k = 0 \text{ for some } k = 1, \ldots, n, \text{ then } x_l = 0 \ \forall \ l \geq k \}.
$$

Then:

$$
\mathbb{E}^P h(X_{t_1}, \ldots, X_{t_n}) = \mathbb{E}^Q g \left( \frac{1}{X_{t_1}}, \ldots, \frac{1}{X_{t_n}} \right) - \sum_{k=0}^{n-1} \mathbb{E}^Q \left( \mathbb{I}_{\{t_k < \tau^X \leq t_{k+1}\}} \cdot \eta_k(X^k) \right),
$$

where we set $t_0 = 0$ and $X^k = \left( \frac{1}{X_{t_1}}, \ldots, \frac{1}{X_{t_k}} \right)$ for $k = 1, \ldots, n - 1$, $X^0 \equiv 0$.

In particular, if $\eta_k(\cdot) \equiv c_k$, $k = 1, \ldots, n - 1$, are constant, then:

$$
\mathbb{E}^P h(X_{t_1}, \ldots, X_{t_n}) = \mathbb{E}^Q g \left( \frac{1}{X_{t_1}}, \ldots, \frac{1}{X_{t_n}} \right) - \sum_{k=0}^{n-1} c_k \cdot \mathbb{Q} \left( t_k < \tau^X \leq t_{k+1} \right).
$$

Proof. First note that

$$
\mathbb{I}_{\{\tau^X > t_n\}} = \mathbb{I}_{\{\tau^X > t_1\}} \mathbb{I}_{\{\tau^X > t_2\}} \cdots \mathbb{I}_{\{\tau^X > t_{n-1}\}} \mathbb{I}_{\{\tau^X > t_n\}}.
$$

Using the change of measure $dP|_{\mathcal{F}_{t_n}} = \frac{1}{X_{t_n}} dQ|_{\mathcal{F}_{t_n}}$ on $\{\tau^X > t_n\}$ we deduce

$$
\mathbb{E}^P h(X) = \mathbb{E}^Q \left( g \left( \frac{1}{X} \right) \mathbb{I}_{\{\tau^X > t_n\}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X} \right) \mathbb{I}_{\{\tau^X > t_1\}} \cdots \mathbb{I}_{\{\tau^X > t_{n-1}\}} \right) =
$$

$$
\mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_1\}} \mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_2\}} \cdots \mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_{n-1}\}} g \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-1}}} \right) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right) \cdots \mathbb{I}_{\mathcal{F}_{t_1}} \right).
$$

Because on $\{\tau^X > t_{n-1}\}$ we have

$$
\mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_n\}} g \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-1}}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-1}}} \right) - \mathbb{E}^Q \left( \mathbb{I}_{\{t_{n-1} < \tau^X \leq t_n\}} \eta_{n-1} (X^{n-1}) \mathbb{I}_{\mathcal{F}_{t_{n-1}}} \right),
$$

it follows that

$$
\mathbb{E}^P h(X) = \mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_1\}} \mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_2\}} \cdots \mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_{n-2}\}} g \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \cdots \mathbb{I}_{\mathcal{F}_{t_1}} \right) \right) \right) - \mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_1\}} \mathbb{I}_{\{\tau^X > t_2\}} \cdots \mathbb{I}_{\{\tau^X > t_{n-1}\}} \eta_{n-1} (X^{n-1}) \mathbb{I}_{\mathcal{F}_{t_{n-1}}} \right).
$$

Similarly, on $\{\tau^X > t_{n-2}\}$ we have

$$
\mathbb{E}^Q \left( \mathbb{I}_{\{\tau^X > t_{n-1}\}} \tilde{g} \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right) = \mathbb{E}^Q \left( \tilde{g} \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right) - \mathbb{E}^Q \left( \mathbb{I}_{\{t_{n-2} < \tau^X \leq t_{n-1}\}} \eta_{n-2} (X^{n-2}) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right),
$$

$$
\mathbb{Q} \left( \mathbb{I}_{\{\tau^X > t_{n-1}\}} \tilde{g} \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right) = \mathbb{Q} \left( \tilde{g} \left( \frac{1}{X} \right) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right) - \mathbb{Q} \left( \mathbb{I}_{\{t_{n-2} < \tau^X \leq t_{n-1}\}} \eta_{n-2} (X^{n-2}) \mathbb{I}_{\mathcal{F}_{t_{n-2}}} \right).
$$
and we deduce that
\[ \mathbb{E}^P h(X) = \mathbb{E}^Q \left( \prod_{\tau X > t_1} \mathbb{E}^Q \left( \prod_{\tau X > t_2} \cdots \mathbb{E}^Q \left( \prod_{\tau X > t_{n-2}} \mathbb{E}^Q \left( \prod_{\tau X > t_{n-1}} \mathbb{E}^Q \left( 1 \right) \right) \right) \right) \right) - \mathbb{E}^Q \left( \prod_{\tau X \leq t_{n-1}} \mathbb{E}^Q \left( \prod_{\tau X \leq t_{n-2}} \cdots \mathbb{E}^Q \left( \prod_{\tau X \leq t_1} \mathbb{E}^Q \left( 1 \right) \right) \right) \right) \]

Iterating this procedure results in
\[ \mathbb{E}^P h(X) = \mathbb{E}^Q \left( 1 \right) - \sum_{k=1}^{n-1} \mathbb{E}^Q \left( \prod_{\tau X \leq t_{k+1}} \mathbb{E}^Q \left( 1 \right) \right) \]

**Remark 3.4.** The second term in the above decompositions (3) and (4) with the minus sign in front will be called the **default term** in the following. This is motivated by the following observation:
\[ \gamma_X(t, T) = X_t - \mathbb{E}^P (X_T | \mathcal{F}_t) = X_t - X_t \cdot Q(\tau^X > T | \mathcal{F}_t) = X_t \cdot Q(\tau^X \leq T | \mathcal{F}_t). \]

Taking expectations under \( P \) yields
\[ \mathbb{E}^P \gamma_X(t, T) = \mathbb{E}^P \left( X_t \cdot Q(\tau^X \leq T | \mathcal{F}_t) \right) = \mathbb{E}^Q \left( \prod_{\tau X > t} Q(\tau^X \leq T | \mathcal{F}_t) \right) = Q(t < \tau^X \leq T). \]

Thus, the default term is directly related to the expected bubble of the underlying.

The convergence conditions that must be fulfilled in Theorem 3.3 may seem to be rather strict. However, below we give a few examples of options which satisfy those conditions.

**Example 3.5.** Let us consider a modified call option with maturity \( T \) and strike \( K \), where the holder has the option to reset the strike value to the current stock price at certain points in time \( t_1 < t_2 < \cdots < t_n < T \), i.e. the payoff profile of the option is given by
\[ H(X) = (X_T - \min(K, X_{t_1}, X_{t_2}, \ldots, X_{t_n}))^+. \]

With the notation in Theorem 3.3 it follows that
\[ \eta_0 = \eta_1 = \cdots = \eta_{n-1} = 1 \]

and the option value can be decomposed as
\[ \mathbb{E}^P h(X) = \mathbb{E}^Q \left( 1 - \frac{1}{X_T} \cdot \min(K, X_{t_1}, \ldots, X_{t_n}) \right)^+ - \sum_{k=0}^{n-1} Q(t_k < \tau^X \leq t_{k+1}) \]
\[ = \mathbb{E}^Q \left( 1 - \frac{1}{X_T} \cdot \min(K, X_{t_1}, \ldots, X_{t_n}) \right)^+ - \gamma_X(0, T). \]

Therefore, this modified call option has the same default as the normal call option, cf. (14) in [28].
Example 3.6. Let us consider a call option on the ratio of the stock price at times $S$ and $T \geq S$ with strike $K \in \mathbb{R}_+$, i.e.

$$h(X) = \left( \frac{X_T}{X_S} - K \right)^+$$

for $S < T \in \mathbb{R}_+$. In this case

$$\eta_0 = 0, \quad \eta_1(y) = y$$

and the decomposition of the option value is given by

$$\mathbb{E}^P h(X) = \mathbb{E}^Q \left( \frac{1}{X_S} - \frac{K}{X_T} \right)^+ - \mathbb{E}^Q \left( \mathbb{1}_{\{S < \tau^X \leq T\}} \frac{1}{X_S} \right).$$

Example 3.7. A chooser option with maturity $T$ and strike $K$ entitles the holder to decide at time $S < T$, whether the option is a call or a put. He will choose the call, if its value is at least as high as the value of the put option with strike $K$ and maturity $T$ at time $S$. However, in the presence of asset price bubbles, i.e. when the underlying is a strict local martingale, put-call-parity does not hold, but instead we have

$$\mathbb{E}^P ((X_T - K)^+ | \mathcal{F}_S) - \mathbb{E}^P ((K - X_T)^+ | \mathcal{F}_S) = \mathbb{E}^P (X_T | \mathcal{F}_S) - K.$$

Therefore, the payoff of the chooser option equals

$$h(X_S, X_T) = (X_T - K)^+ \mathbb{1}_{\{\mathbb{E}^P (X_T | \mathcal{F}_S) \geq K\}} + (K - X_T)^+ \mathbb{1}_{\{\mathbb{E}^P (X_T | \mathcal{F}_S) < K\}}.$$

Let us assume that $X$ is Markovian. Then we can express $\mathbb{E}^P (X_T | \mathcal{F}_S)$ as a function of $X_S$, say $\mathbb{E}^P (X_T | \mathcal{F}_S) = m(X_S)$, and the limits defined in Theorem 3.3 exist, if $m$ is monotone for large values, and equal

$$\eta_1(y) = \mathbb{1}_{\{m\left(\frac{1}{y}\right) \geq K\}}, \quad \eta_0 = \lim_{x \to \infty} \mathbb{1}_{\{m(x) \geq K\}}.$$

Thus, the value of the chooser option can be decomposed as

$$\mathbb{E}^P h(X_S, X_T) = \mathbb{E}^Q \left( \frac{h(X_S, X_T)}{X_T} \right) - \mathbb{Q} \left( m(X_S) \geq K, \ S < \tau^X \leq T \right) - \lim_{x \to \infty} \mathbb{1}_{\{m(x) \geq K\}} \mathbb{Q} (\tau^X \leq S).$$

If $X$ is the inverse of a BES(3)-process under $P$, it is calculated in subsection 2.2.2 in [5] that

$$m(X_s) = \mathbb{E}^P (X_T | X_S) = X_S \left( 1 - 2\Phi \left( -\frac{1}{X_S \sqrt{T - S}} \right) \right).$$

Therefore,

$$\lim_{x \to \infty} m(x) = \lim_{x \to \infty} \mathbb{E}^P (X_T | X_S = x) = \lim_{x \to \infty} 2\varphi \left( -\frac{1}{x \sqrt{T - S}} \right) \frac{1}{\sqrt{T - S}} = \frac{\sqrt{2}}{\sqrt{\pi(T - S)}}$$

and

$$\eta_1(y) = \mathbb{1}_{\left\{ \frac{1}{y} \left( 1 - 2\Phi \left( -\frac{\sqrt{y}}{\sqrt{T - S}} \right) \right) \geq K \right\}}, \quad \eta_0 = \mathbb{1}_{\left\{ \frac{\sqrt{2}}{\sqrt{\pi(T - S)}} > K \right\}}.$$

In the following we give another extension of Proposition 7 in [28] to Barrier options, i.e. we allow the options to be knocked-in or knocked-out by passing some barrier.
Theorem 3.8. Consider any non-negative Borel-measurable function $h : \mathbb{R}_{++} \to \mathbb{R}_+$ and define $g(x) = x \cdot h \left( \frac{1}{x} \right)$ for $x > 0$. Suppose that $\lim_{x \to 0} g(x) =: \eta < \infty$ exists and denote by $\bar{g} : \mathbb{R}_+ \to \mathbb{R}_+$ the extension of $g$ with $g(0) = \eta$. Define $\hat{m}_T^X = \min_{t \leq T} X_t$, $m_T^X = \max_{t \leq T} X_t$, and $\tau_a^X := \inf\{ t \geq 0 : X_t > a \}$, $T_a^X := \inf\{ t \geq 0 : X_t \leq a \}$ for $a \in \mathbb{R}_+$. Then for any bounded stopping time $T$ and for any real numbers $D \leq 1$ and $F \geq 1$:

$$
\begin{align*}
\text{(DI)} & \quad \mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ \hat{m}_T^X \leq D \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \hat{m}_T^X \leq D \}} \right) - \eta \cdot \mathbb{Q} \left( T_D^X < \tau^X \leq T \right) \\
\text{(DO)} & \quad \mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ m_T^X \geq D \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ m_T^X \geq D \}} \right) - \eta \cdot \mathbb{Q} \left( T_D^X = \infty, \tau^X \leq T \right) \\
\text{(UI)} & \quad \mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ m_T^X \geq F \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ m_T^X \geq F \}} \right) - \eta \cdot \mathbb{Q} \left( \tau^X \leq T \right) \\
\text{(UO)} & \quad \mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ m_T^X \leq F \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ m_T^X \leq F \}} \right)
\end{align*}
$$

Before proving the theorem we remark that the result is intuitively reasonable in that the default only plays a role if the option is active. Especially note that the default term for Up-and-out options is equal to zero, a result which is true for all options with bounded payoffs (cf. Corollary 1 in [28]).

Proof. Keeping in mind that $D \leq 1$ and $F \geq 1$, it follows from the absolute continuity relationship between $P$ and $Q$ that

$$
\begin{align*}
\mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ \hat{m}_T^X \leq D \}} \right) & = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \tau^X > T, \hat{m}_T^X \leq D \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \tau^X > T, T_D^X \}} \right) \\
& = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \hat{m}_T^X \leq D \}} \right) - \eta \cdot \mathbb{Q} \left( T_D^X \leq T, \tau^X \leq T \right) \\
& = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \hat{m}_T^X \leq D \}} \right) - \eta \cdot \mathbb{Q} \left( T_D^X < \tau^X \leq T \right).
\end{align*}
$$

Similarly,

$$
\begin{align*}
\mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ \hat{m}_T^X \geq D \}} \right) & = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \tau^X > T, \hat{m}_T^X \geq D \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \tau^X > T, T_D^X \}} \right) \\
& = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \hat{m}_T^X \geq D \}} \right) - \eta \cdot \mathbb{Q} \left( \tau^X \leq T < T_D^X \right) \\
& = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \hat{m}_T^X \geq D \}} \right) - \eta \cdot \mathbb{Q} \left( \tau^X \leq T, T_D^X = \infty \right).
\end{align*}
$$

Continuing,

$$
\begin{align*}
\mathbb{E}^P \left( h(X_T) \mathbb{1}_{\{ m_T^X \geq F \}} \right) & = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \tau^X > T, m_T^X \geq F \}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ \tau^X > T, \tau_F^X \}} \right) \\
& = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ m_T^X \geq F \}} \right) - \eta \cdot \mathbb{Q} \left( \tau_F^X \leq T, \tau^X \leq T \right) \\
& = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{1}_{\{ m_T^X \geq F \}} \right) - \eta \cdot \mathbb{Q} \left( \tau^X \leq T \right).
\end{align*}
$$
Finally,

\[
\mathbb{E}^P \left( h(X_T) \mathbb{I}_{\{m_X^T \leq F\}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{I}_{\{\tau^X > T, m_X^T \leq F\}} \right) = \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{I}_{\{\tau^X > T, \tau^F > T\}} \right)
\]

\[
= \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{I}_{\{m_X^T \leq F\}} \right) - \eta \cdot \mathbb{Q}(\tau^X \leq T < \tau^F)
\]

\[
= \mathbb{E}^Q \left( g \left( \frac{1}{X_T} \right) \mathbb{I}_{\{m_X^T \leq F\}} \right).
\]

All the formulas have been proved. □

**Remark 3.9.** Above we used the risk-neutral pricing approach to calculate the value of some options written on stocks which may have an asset price bubbles, as suggested by the first fundamental theorem of asset pricing. The derived decompositions show that there is an important difference in the option value depending on whether the underlying is a strict local or a true martingale under the risk-neutral measure, which is reflected in the default term. Even though we do not create arbitrage opportunities when pricing options by their fundamental values calculated above, several authors have suggested to “correct” the option price to account for the strictness of the local martingale, cf. e.g. [17, 18, 19, 24]. While the authors of [17, 18, 19] work under the additional No Dominance assumption which is strictly stronger than NFLVR, Madan and Yor suggest in [24] the following pricing formulas for European and American call options written on (continuous) \( X \) with strike \( K \) and maturity \( T \):

\[
C^\text{strict}_E(K,T) := \lim_{n \to \infty} \mathbb{E}^P(X_{T \land \sigma_n} - K)^+,
\]

\[
C^\text{strict}_A(K,T) := \sup_{\sigma \in T_{0,T}} \lim_{n \to \infty} \mathbb{E}^P(X_{\sigma \land \sigma_n} - K)^+
\]

for some localizing sequence \((\sigma_n)_{n \in \mathbb{N}}\) of the strict local martingale \( X \). It is proven in [24] that this definition is independent of the chosen localizing sequence and that \( C^\text{strict}_E = C^\text{strict}_A \). However, a generalization of this definition to any other option \( h(\cdot) \) on \( X \) with maturity \( T \) is problematic: the independence of the chosen localizing sequence \((\sigma_n)_{n \in \mathbb{N}}\) is not true in general, so one may have to choose \( \sigma_n = \tau^X_n \) as defined above. Moreover, in general \( \lim_{n \to \infty} \mathbb{E}^P h(X^{\sigma_n}) \) may not even be well-defined and equal to \( \mathbb{E}^P h(X_T) \), when \( X \) is a true martingale. Since in this case there are no asset price bubbles, it does not seem correct to trade the option for a price which differs from its fundamental value. Therefore, in the case where we have a decomposition of the fundamental option value as above or more generally as proven in Theorem 3.3, this suggests that the most sensible approach to correct the option value for bubbles in the underlying is to set the default term equal to zero. Equivalently, we can also set \( \tau^X \) equal to infinity under the measure \( Q \). This even gives a way of correcting the option value for stock price bubbles in the general case, where a decomposition formula may not be available. By doing so we basically treat the price process as if it was a true martingale. However, we want to emphasize that it is not necessary to correct the price at all, since the fundamental value gives an arbitrage-free price.
4. Relationship Between $P$ and $Q$

In the following we study the relationship between the original measure $P$ and the measure $Q$ in more detail. To avoid an overload of notation in what follows, we introduce the following convention which will be valid throughout:

\[ X \text{ is assumed to be a càdlàg strictly positive local martingale on the probability space } (\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, P), \]
\[ \text{whose filtration is the right-continuous augmentation of a standard system and } \mathcal{F} = \bigvee_{t \geq 0} \mathcal{F}_t. \]
\[ \text{We set } \tau^X_n = \inf\{t \geq 0 | X_t > n\} \land n \text{ for all } n \in \mathbb{N} \text{ and } \tau^X = \lim_{n \to \infty} \tau^X_n. \]
\[ \text{Furthermore, we denote by } Q \text{ any extension to } (\Omega, \mathcal{F}) \text{ of the measure associated with } X, \text{ defined in Theorem 1.8.} \]

**Lemma 4.1.** Set \( X = \tilde{X} \), i.e. \( X_t = \infty \) on \( \{t \geq \tau^X\} \). Then, \( Q(X_{\infty} = \infty) = 1 \iff P(X_{\infty} = 0) = 1 \).

**Proof.** Since \( X \) is a \( P \)-supermartingale and \( \frac{1}{X} \) a \( Q \)-martingale, both converge and therefore \( X_{\infty} \) is almost surely well-defined under both measures.

\( \Leftarrow: \) Assume that \( P(X_{\infty} = 0) = 1 \). Because \( 1/X \) is a \( Q \)-martingale, we have by Fatou’s lemma for all \( u > 0 \),
\[
\mathbb{E}^Q \left( \frac{1}{X_{\infty}} I_{\{\tau^X > t, X_t > u\}} \right) \leq \liminf_{n \to \infty} \mathbb{E}^Q \left( \frac{1}{X_{t+n}} I_{\{\tau^X > t, X_t > u\}} \right) = \mathbb{E}^Q \left( \frac{1}{X_t} I_{\{\tau^X > t, X_t > u\}} \right) = P(X_t > u).
\]

By dominated convergence for \( t \to \infty \),
\[
\mathbb{E}^Q \left( \frac{1}{X_{\infty}} I_{\{\tau^X = \infty, X_{\infty} > u\}} \right) \leq P(X_{\infty} \geq u) = 0 \quad \forall u > 0.
\]

This implies that
\[
\mathbb{E}^Q \left( \frac{1}{X_{\infty}} I_{\{\tau^X = \infty\}} \right) = 0.
\]

Since \( \frac{1}{X} \) is a \( Q \)-martingale,
\[
\mathbb{E}^Q \left( \frac{1}{X_{\infty}} \right) \leq \mathbb{E}^Q \left( \frac{1}{X_t} \right) = 1.
\]

Thus, \( Q(X_{\infty} = 0) = 0 \) and
\[
\mathbb{E}^Q \left( \frac{1}{X_{\infty}} I_{\{\tau^X = \infty\}} \right) = 0 \iff \frac{1}{X_{\infty}} I_{\{\tau^X = \infty\}} = 0 \text{ } Q\text{-a.s.}
\]

Since \( \frac{1}{X_{\infty}} I_{\{\tau^X < \infty\}} = 0 \), it follows that \( \frac{1}{X_{\infty}} = 0 \text{ } Q\text{-almost surely.}
\]

\( \Rightarrow: \) Assume that \( Q(X_{\infty} = \infty) = 1 \). Because \( X \) is a \( P \)-supermartingale, we have
\[
\mathbb{E}^P X_{\infty} \leq \mathbb{E}^P X_t \leq 1
\]
and
\[
\mathbb{E}^P \left( X_{\infty} I_{\{X_t < k\}} \right) \leq \mathbb{E}^P \left( X_t I_{\{X_t < k\}} \right) = Q(t < \tau^X, X_t < k) = Q(X_t < k) \quad \forall k \geq 0.
\]

For \( t \to \infty \) by dominated convergence then
\[
\mathbb{E}^P \left( X_{\infty} I_{\{X_{\infty} < k\}} \right) \leq Q(X_{\infty} < k) = 0 \quad \forall k \geq 0.
\]
This implies that \( X\infty \mathbb{I}_{\{X_\infty < k\}} = 0 \) \( \mathbb{P} \)-almost surely for all \( k \geq 0 \). Therefore, \( \mathbb{P}(X_\infty \in \{0, \infty\}) = 1 \). Since \( \mathbb{E}^\mathbb{P}(X_\infty) \leq 1 \), it follows that \( \mathbb{P}(X_\infty = \infty) = 0 \) and thus \( X_\infty = 0 \) \( \mathbb{P} \)-almost surely. \( \square \)

Until here we have only considered the behaviour of the local \( \mathbb{P} \)-martingale \( X \) under \( \mathbb{Q} \). But how do other processes change their behaviour, when passing from \( \mathbb{P} \) to \( \mathbb{Q} \)? This question is of particular interest, since we want to apply our result to option pricing, where under the risk-neutral measure all discounted stock price processes turn into local martingales. Let us assume that besides \( X \) do other processes change their behaviour, when passing from \( \mathbb{P} \) to \( \mathbb{Q} \). We shall therefore introduce a localizing process \( \bar{Y} \) under \( \mathbb{Q} \).

**Lemma 4.2.** Let \( Y \) be a strictly positive càdlàg local \( \mathbb{P} \)-martingale. Then: \( \mathbb{Q}(\tau^Y \leq \tau^X) = 1 \).

**Proof.**

\[
\mathbb{Q}(\tau^Y < \tau^X) = \lim_{n \to \infty} \mathbb{Q}(\tau^Y < \tau^X_n) = \lim_{n \to \infty} \mathbb{E}^\mathbb{P}\left( X_{\tau^X_n} \mathbb{I}_{\{\tau^Y < \tau^X_n\}} \right) = 0.
\]

\( \square \)

Moreover, we introduce condition (A): \( \mathbb{Q}(Y_{\tau^X} = \infty) = 0 \).

It can be shown that (A) is always fulfilled, if \( X \) and \( Y \) are independent or if \( X \) is a true martingale. However, in general it is hard to check (A), since it requires some knowledge of the joint distribution of \( \tau^X_n \) and \( \tau^Y_m \) for \( n, m \) large.

If \( X \) and \( Y \) are assumed to be càdlàg processes under \( \mathbb{P} \), they are also almost surely càdlàg under \( \mathbb{Q} \) before time \( \tau^X \) because \( \mathbb{P} \) and \( \mathbb{Q} \) are equivalent on every \( \mathcal{F}_{\tau^X} \). Furthermore, since \( \frac{1}{X} \) is a \( \mathbb{Q} \)-martingale, it does not explode and therefore \( X_{\tau^X -} \neq 0 \) and \( X_t \neq 0 \) \( \mathbb{Q} \)-almost surely for all \( t \geq 0 \). Thus, the process \( Z := \frac{Y}{X} \) does also have almost surely càdlàg paths before time \( \tau^X \). Since from time \( \tau^X \) on everything is constant, the only crucial question is whether \( Z = \frac{Y}{X} \) has a left-limit at \( \tau^X \).

**Lemma 4.3.** Let \( Y \) be a non-negative local \( \mathbb{P} \)-martingale. Then \( Z_t := \left( \frac{Y_t}{X_t} \right)_{0 \leq t < \tau^X} \) is a local martingale on \( (\Omega, \mathcal{F}_{\tau^X -}, (\mathcal{F}_{t \wedge \tau^X -})_{t \geq 0}, \mathbb{Q}) \). Furthermore, setting \( Z_t := \bar{Z}_t \) and \( X_t = \infty \) on \( \{t \geq \tau^X\} \) is the unique way to define \( Z \) and \( X \) after time \( \tau^X \) such that \( \frac{1}{X} \) and \( Z \) remain non-negative càdlàg local martingales on \( [0, \infty) \) for all possible extensions of the measure \( \mathbb{Q} \) from \( \mathcal{F}_{\tau^X -} \) to \( \mathcal{F} = \bigvee_{t \geq 0} \mathcal{F}_t \).

**Proof.** First, we show that \( Z = \frac{Y}{X} \) is a local \( \mathbb{Q} \)-martingale on \( \bigcup_{n \in \mathbb{N}} [0, \tau^X_n] \) with localizing sequence \( (\tau^X_n \wedge \tau^X_m)_{n,m \in \mathbb{N}} \). Indeed, for \( t \geq s \) we have for all \( n \in \mathbb{N} \)

\[
\mathbb{E}^\mathbb{Q}\left( Z_{t \wedge \tau^X_n \wedge \tau^X_m} \mid \mathcal{F}_s \right) = \mathbb{E}^\mathbb{Q}\left( \frac{Y_{t \wedge \tau^X_n \wedge \tau^X_m}}{X_{t \wedge \tau^X_n \wedge \tau^X_m}} \mid \mathcal{F}_s \right) = \mathbb{E}^\mathbb{P}\left( \frac{Y_{s \wedge \tau^X_n \wedge \tau^X_m}}{X_{s \wedge \tau^X_n \wedge \tau^X_m}} \mid \mathcal{F}_s \right) = \frac{Y_{s \wedge \tau^X_n \wedge \tau^X_m}}{X_{s \wedge \tau^X_n \wedge \tau^X_m}} = Z_{s \wedge \tau^X_n \wedge \tau^X_m}.
\]
and by the first part of Lemma 4.2, we know that $\tau^X_n \wedge \tau^Y_n \to \tau^X$ Q-almost surely. Since $Z$ is a non-negative local supermartingale on $\bigcup_{n \in \mathbb{N}} [0, \tau^X_n)$, we can apply Fatou’s lemma twice with $s \leq t$:

$$
\hat{Z}_s = \liminf_{u \to \tau^X_n, u < \tau^X_n, u \in \mathbb{Q}} Z_{s \wedge u} = \liminf_{u \to \tau^X_n, u < \tau^X_n, u \in \mathbb{Q}^n \to \infty} \lim_{n \to \infty} Z_{s \wedge u \wedge \tau^X_n \wedge \tau^Y_n} \\
\geq \liminf_{u \to \tau^X_n, u < \tau^X_n, u \in \mathbb{Q}^n \to \infty} \mathbb{E}^Q \left( Z_{t \wedge u \wedge \tau^X_n \wedge \tau^Y_n} \mid \mathcal{F}_s \right) \geq \liminf_{u \to \tau^X_n, u < \tau^X_n, u \in \mathbb{Q}} \mathbb{E}^Q (Z_{t \wedge u} \mid \mathcal{F}_s) \\
\geq \mathbb{E}^Q \left( \liminf_{u \to \tau^X_n, u < \tau^X_n, u \in \mathbb{Q}} Z_{t \wedge u} \mid \mathcal{F}_s \right) = \mathbb{E}^Q (\hat{Z}_t \mid \mathcal{F}_s),
$$

where the second inequality is due to the fact that $\mathbb{E}^Q (Z_{t \wedge u \wedge \tau^X_n \wedge \tau^Y_n} \mid \mathcal{F}_s) \geq \mathbb{E}^Q (Z_{t \wedge u} \mid \mathcal{F}_s)$ by the supermartingale property. By the convergence theorem for positive supermartingales, we conclude that $\hat{Z}_{\tau^{X,-}} = Z_{\tau^{X,-}}$ exists Q-almost surely in $\mathbb{R}_+$. To see that $\hat{Z}$ is indeed a local martingale and not only a supermartingale, we show that $\hat{Z}^{\tau^X}_{n}$ is a uniformly integrable martingale for all $n \in \mathbb{N}$, where $\tau^Z_n = \inf\{ t \geq 0 \mid Z_t > n \} \wedge n$. Since $\hat{Z}$ is a non-negative supermartingale, it is sufficient to prove that the expectation of $\hat{Z}^{\tau^X}_{n}$ is constant:

$$
\mathbb{E}^Q \hat{Z}^{\tau^X}_{m} = \mathbb{E}^Q \left( \hat{Z}^{\tau^X}_{m} \mid \tau^{\tau^X}_{\tau^X \wedge \tau^Y} + \hat{Z}^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) = \lim_{m \to \infty} \mathbb{E}^Q \left( Z^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) + \mathbb{E}^Q \left( \hat{Z}^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) \\
= \lim_{m \to \infty} \mathbb{E}^Q \left( Z^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) + \mathbb{E}^Q \left( \lim_{m \to \infty} Z^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) \\
= \lim_{m \to \infty} \mathbb{E}^Q \left( Z^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) + \lim_{m \to \infty} \mathbb{E}^Q \left( Z^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) \\
= \lim_{m \to \infty} \mathbb{E}^Q \left( Z^{\tau^X}_{m} \tau^{\tau^X}_{\tau^X \wedge \tau^Y} \right) = Z_0.
$$

To prove the uniqueness of the extension of $Z$ for all possible extensions of $Q$ to $\mathcal{F}$, define for all $n \in \mathbb{N}$ $\tau^{Z}_{n} = \inf\{ t \geq 0 \mid Z_t > n \}$, where $Z$ is an arbitrary càdlàg extension of $(Z_t)_{t<\tau^X}$. Then $(\tau^{Z}_{n})_{n \in \mathbb{N}}$ is a localizing sequence for $Z$ for all possible extensions of $Q$. Fix one of these extensions and call it $Q^0$. We have

$$
\mathbb{E}^{Q^0} (Z^{\tau^X}_{t} \mid \mathcal{F}_s) = Z^{\tau^X}_{s} \quad \forall \ n \in \mathbb{N}.
$$

Now for fix $n \in \mathbb{N}$ define the new measure $Q^n$ on $\mathcal{F}$ via

$$
\frac{dQ^n}{dQ^0} = \frac{Z^{\tau^X}_{\tau^{Z}_{n}}}{Z^{\tau^X}_{\tau^{Z}_{n}}}.
$$

Note that $Q^n$ is also an extension of $Q$ from $\mathcal{F}_{\tau^{X,-}}$ to $\mathcal{F}$. Furthermore, for all $\varepsilon \geq 0$, we have

$$
\mathbb{E}^{Q^n} \left( Z^{\tau^X}_{\tau^{X} + \varepsilon} \mid \mathcal{F}_{\tau^{X} -} \right) = \mathbb{E}^{Q^n} \left( \frac{Z^{\tau^X}_{\tau^{Z}_{n}}}{Z^{\tau^X}_{\tau^{Z}_{n}}} \cdot Z^{\tau^X}_{\tau^{X} + \varepsilon} \mid \mathcal{F}_{\tau^{X} -} \right) = \mathbb{E}^{Q^n} \left( \left( \frac{Z^{\tau^X}_{\tau^{X} + \varepsilon}}{Z^{\tau^X}_{\tau^{Z}_{n}}} \right)^2 \mid \mathcal{F}_{\tau^{X} -} \right),
$$

because $\hat{Z}^{\tau^X}_{n}$ must also be a uniformly integrable martingale under $Q^n$. Therefore, $\hat{Z}^{\tau^X}_{n}$ and $(\hat{Z}^{\tau^X}_{n})^2$ are both $Q^0$-martingales, which implies that $Z_{\varepsilon + \tau^X} = Z_{\tau^{X} -}$ for all $\varepsilon \geq 0$. Thus, $Z \equiv \hat{Z}$ is uniquely determined.

As usual to simplify notation we will identify $Z$ with the process $\hat{Z}$ in the following.
Remark 4.4.

- Note that if (A) is satisfied, then $Z_{\tau^X} = 0$ on $\{\tau^X < \infty\}$ $Q$-almost surely.
- Even though we proved that $Z_{\tau^X}$ exists $Q$-almost surely and also $X_{\tau^X}$ is well-defined, this does not allow us to infer any conclusions about the set $\{Y_{\tau^X} \text{ exists in } \mathbb{R}_+\}$ in general.
- For our purposes it is sufficient that local $Q$-martingales are càdlàg almost everywhere, since we are only interested in pricing and do not deal with an uncountable number of processes. One should however have in mind that in order to have everywhere regular paths some kind of augmentation is needed, cf. [21].

Remark 4.5. If $\Omega = C_{\infty}(\mathbb{R}_+, \mathbb{R}^2)$ is the pathspace introduced in Lemma 1.5 and $(X, Y)$ is the coordinate process as well as $(\mathcal{F}_t)_{t \geq 0}$ the canonical filtration generated by $(X, Y)$, then we extend $Q$ to $\mathcal{F} = \bigvee_{t \geq 0} \mathcal{F}_t$ such that

$$Q(\omega_1(t) = \infty, \omega_2(t) = \omega_2(\tau^X) \forall t \geq \tau^X) = 1.$$ 

Lemma 4.6.

(1) If $X$ is a $P$-martingale, then $Z := \frac{Y}{X}$ is a strict local $Q$-martingale if and only if $Y$ is a strict local $P$-martingale.

(2) Assume that $X$ is a strict local $P$-martingale. Then:
   (a) If $Y$ is a $P$-martingale, then $Z$ is a $Q$-martingale and $Z_{\tau^X} = 0$ on $\{\tau^X < \infty\}$.
   (b) If $Z$ is a strict local $Q$-martingale, then $Y$ is a strict local $P$-martingale.
   (c) If $Z$ is a $Q$-martingale and if (A) holds, then $Y$ is a $P$-martingale.
   (d) If $Y$ is a strict local $P$-martingale and if (A) holds, then $Z$ is a strict local $Q$-martingale.

Proof.

(1) This is obvious, because $Q$ and $P$ are locally equivalent, if $X$ is a true $P$-martingale.

(2) First note that

$$E^P Y_0 = E^Q Z_0 \geq E^Q Z_t = E^Q \left( Z_t \mathbb{I}_{\{t < \tau^X\}} \right) + E^Q \left( Z_t \mathbb{I}_{\{t \geq \tau^X\}} \right) = E^Q \left( \frac{Y_t}{X_t} \mathbb{I}_{\{t < \tau^X\}} \right) + E^Q \left( Z_{\tau^X} \mathbb{I}_{\{t \geq \tau^X\}} \right) = E^P Y_t + E^Q \left( Z_{\tau^X} \mathbb{I}_{\{t \geq \tau^X\}} \right).$$

(a) Since $Y$ is a positive local $P$-martingale, we have:

$Y$ is a true $P$-martingale $\iff E^P Y_t = E^P Y_0$ for all $t \geq 0$

$$\iff E^Q Z_t = E^Q Z_0 \text{ for all } t \geq 0, \text{ } Z_{\tau^X} \mathbb{I}_{\{\tau^X < \infty\}} = 0 \text{ } Q\text{-almost surely.}$$

(b) Obviously, if $E^Q Z_t < E^Q Z_0$ for some $t \geq 0$, the above inequality is strict and thus also $E^P Y_t < E^P Y_0$.

(c) If (A) holds, $Z_{\tau^X} = 0$ on $\{\tau^X < \infty\}$ $Q$-almost surely by Remark 4.5. Therefore, since $Z$ is a $Q$-martingale, the above inequality turns into an equality and $Y$ is a true $P$-martingale.
(d) Assume that $Y$ is a strict local $P$-martingale, i.e. $E^P Y_t < E^P Y_0$ for some $t > 0$. Furthermore assume that (A) holds, which implies that $Z_{r_X} = 0$ $Q$-almost surely on $\{r_X < \infty\}$. Then:

$$E^Q Z_t = E^P Y_t + E^Q \left( Z_{r_X} I_{\{t \geq r_X\}} \right) = E^P Y_t < E^P Y_0 = E^Q Z_0,$$

i.e. $Z$ is a strict local $Q$-martingale.

□

Example 4.7. (Continuation of Example 2.1.3)

For the following example we work on the pathspace $C_\infty([0, \infty), \mathbb{R}^2)$ with $(X, Y)$ denoting the coordinate process and $(F_t)_{t \geq 0}$ being the right-continuous augmentation of the canonical filtration generated by the coordinate process. Remember from example 2.1.3 that for $\sigma(x)$ locally bounded and bounded away from zero for $x > 0$, $\sigma(0) = 0$, the local $P$-martingale

$$dX_t = \sigma(X_t) dW_t, \quad X_0 = 1,$$

is strictly positive whenever

$$\int_0^1 \frac{x}{\sigma^2(x)} dx = \infty,$$

and under $Q$ with $\frac{dP}{dQ} \big|_{F_1} = \frac{1}{X_1}$ the inverse process is a true martingale with decomposition

$$d \left( \frac{1}{X_t} \right) = -\frac{\sigma(X_t)}{X_t^2} dW_t^Q = \bar{\sigma} \left( \frac{1}{X_t} \right) dW_t^Q$$

for a $Q$-Brownian motion $W^Q$ and $\bar{\sigma}(y) := -y^2 \cdot \sigma \left( \frac{1}{y} \right)$.

Now let us assume that $Y$ is also a local martingale under $P$ with dynamics

$$dY_t = \gamma(Y_t) dB_t,$$

where $\gamma$ fulfills the same assumptions as $\sigma$ and $B$ is another $P$-Brownian motion such that $d\langle B, W \rangle_t = \rho dt$. Then $\frac{Y}{X}$ is a $Q$-local martingale with decomposition

$$d \left( \frac{Y_t}{X_t} \right) = \frac{\gamma(Y_t)}{X_t} dB_t^Q + Y_t \bar{\sigma} \left( \frac{1}{X_t} \right) dW_t^Q,$$

where $B^Q$ is a $Q$-BM such that $d\langle B^Q, W^Q \rangle_t = \rho dt$.

5. Applications II: Last passage times formulas

In Section 3 we have seen how one can determine the influence bubbles have on option pricing formulas through a decomposition of the option value into a “normal” term and a default term (cf. Theorems 3.3, 3.8). However this approach only works well for options written on one underlying. It is rather difficult to give a universal way of how to determine the influence of asset price bubbles on the valuation of more complicated options and we will not do this here in all generality. Instead, we will do the analysis for a special example, the so called exchange option, which allows us to connect results about last passage times with the change of measure that was defined from
The setup is the same as in the previous subsection, but now we assume that there exists another strictly positive process $Y$ on $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$, which is also a local $\mathbb{P}$-martingale. In the following we will assume that $X$ and $Y$ are continuous. As before we define $Z = \frac{Y}{X}$, which is a local $\mathbb{Q}$-martingale.

5.1. **Exchange option.** With the interpretation of $X$ and $Y$ as two stock price processes and assuming an interest rate of $r = 0$, we can define the price of a European exchange option with strike $K \in \mathbb{R}_+$ (also known as the ratio of notionals) and maturity $T \in \mathbb{R}_+$ as

$$E(K, T) := \mathbb{E}^\mathbb{P}(X_T - KY_T)^+. $$

The corresponding price of the American option is given by

$$A(K, T) := \sup_{\sigma \in \mathcal{T}_{0,T}} \mathbb{E}^\mathbb{P}(X_{\sigma} - KY_{\sigma})^+, $$

where $\mathcal{T}_{0,T}$ is the set of all stopping times $\sigma$, which take values in $[0, T]$. Let us define the last passage time $\rho_K := \sup \{ t \geq 0 | Z_t = \frac{1}{K} \}$, where as usual the supremum of the empty set is equal to zero. In the next theorem the prices of European and American exchange options are expressed in terms of the last passage time $\rho_K$ in the spirit of [31].

**Theorem 5.1.** For all $K, T \geq 0$ the prices of the European and American exchange option are given by

$$E(K, T) = \mathbb{E}^\mathbb{Q}\left((1 - KZ_{\tau_x})^+ \mathbb{I}_{\{\rho_K \leq T < \tau_x\}}\right), \quad A(K, T) = \mathbb{E}^\mathbb{Q}\left((1 - KZ_{\tau_x})^+ \mathbb{I}_{\{\rho_K \leq \tau_x \wedge T\}}\right).$$

**Proof.** Assume $\sigma \in \mathcal{T}_{0,T}$. As seen above, $Z = \frac{Y}{X}$ is a non-negative local $\mathbb{Q}$-martingale, thus a supermartingale, which converges almost surely to $Z_\infty = Z_{\tau_x}$. From Corollary 3.4 in [3] resp. Theorem 2.5 in [31] we have the identity

$$\left(\frac{1}{K} - Z_{\sigma}\right)^+ = \mathbb{E}^\mathbb{Q}\left(\left(\frac{1}{K} - Z_{\tau_x}\right)^+ \mathbb{I}_{\{\rho_K \leq \sigma\}} \bigg| \mathcal{F}_\sigma\right).$$

Multiplying the above equation with the $\mathcal{F}_\sigma$-measurable random variable $K\mathbb{I}_{\{\tau_x > \sigma\}}$ and taking expectations under $\mathbb{Q}$ yields

$$\mathbb{E}^\mathbb{Q}\left((1 - KZ_{\sigma})^+ \mathbb{I}_{\{\tau_x > \sigma\}}\right) = \mathbb{E}^\mathbb{Q}\left((1 - KZ_{\tau_x})^+ \mathbb{I}_{\{\rho_K \leq \sigma \wedge \tau_x\}}\right).$$

Changing the measure via $d\mathbb{P}|_{\mathcal{F}_\sigma} = \frac{1}{X_\sigma} d\mathbb{Q}|_{\mathcal{F}_\sigma}$, we obtain

$$\mathbb{E}^\mathbb{P}(X_{\sigma} - KY_{\sigma})^+ = \mathbb{E}^\mathbb{P}\left(\mathbb{I}_{\{\tau_x > \sigma\}} X_{\sigma} (1 - KZ_{\sigma})^+\right) = \mathbb{E}^\mathbb{Q}\left((1 - KZ_{\tau_x})^+ \mathbb{I}_{\{\rho_K \leq \sigma \wedge \tau_x\}}\right),$$

since $\mathbb{I}_{\{\tau_x > \sigma\}} = 1$ $\mathbb{P}$-almost surely. Taking $\sigma = T$ the formula for the European option is proven.

For the American option value we note that in the proof of Theorem 1.4 in [1] it is shown that

$$A(K, T) = \lim_{n \to \infty} \mathbb{E}^\mathbb{P}\left(Y_{\tau_n \wedge T} \left(\frac{1}{Z_{\tau_n \wedge T}} - K\right)^+\right) = \lim_{n \to \infty} \mathbb{E}^\mathbb{P}\left(X_{\tau_n \wedge T} - KY_{\tau_n \wedge T}\right)^+. $$
Setting $\sigma = \tau_n^X \wedge T$ in equality (6), it follows that
\[
A(K, T) = \lim_{n \to \infty} \mathbb{E}^P \left( X_{\tau_n^X \wedge T} - KY_{\tau_n^X \wedge T}^+ \right) = \lim_{n \to \infty} \mathbb{E}^Q \left( (1 - KZ_{\tau_n^X \wedge T})^+ \mathbb{1}_{(\rho_K \leq \tau_n^X \wedge T < \tau^X)} \right)
\]
\[
= \lim_{n \to \infty} \mathbb{E}^Q \left( (1 - KZ_{\tau_n^X \wedge T})^+ \mathbb{1}_{(\rho_K \leq \tau_n^X \wedge T)} \right) = \mathbb{E}^Q \left( (1 - KZ_{\tau_n^X \wedge T})^+ \mathbb{1}_{(\rho_K \leq \tau_n^X \wedge T)} \right).
\]

\[\square\]

Remark 5.2. If we take $Y \equiv 1$ in the above theorem, we get the formula for the standard European call option expressed as a function of the last passage time of $X$ as it can be found in [32] for the special case of Bessel processes or [22]:
\[
E(K, T) = Q \left( \rho_K \leq T < \tau^X \right).
\]

More generally, formula (7) is always true, if (A) holds.

Remark 5.3. We can also express the price of a barrier exchange option in terms of the last passage time of $Z$ at level $1/K$ as done in Theorem 5.1 for exchange options without barriers. For example, in the case of the Down-and-In exchange option we simply have to multiply equation (5) with the $\mathcal{F}_\sigma$-measurable random variable $\mathbb{1}_{\{\hat{m}_{\sigma}^X \leq D\}}$.

We now analyze a few special cases of Theorem 5.1 in more detail:

1. \textbf{X is a true P-martingale}

   If $X$ is a true $P$-martingale, the price process for $X$ exhibits no asset price bubble. Then, regardless of whether the stock price process $Y$ has an asset price bubble or not, we know that $Q$ is locally equivalent to $P$ and $Q(\tau^X = \infty) = 1$. Therefore
   \[
   E(K, T) = A(K, T) = \mathbb{E}^Q \left( (1 - KZ_{\infty})^+ \mathbb{1}_{(\rho_K \leq T)} \right)
   \]
   and the European and American call option values are equal. For $Y \equiv 1$ this formula is well-known, cf. [31].

2. \textbf{Y is a true P-martingale}

   We recall from Lemma 4.6 that in this case $Z_{\tau^X} = 0$ on $\{\tau^X < \infty\}$ $Q$-almost surely. Denoting $\tau_0^Z = \inf\{t \geq 0 \mid Z_t = 0\}$ this translates into $Q(\tau^X = \tau_0^Z) = 1$, since
   \[
   Q(\tau_0^Z < \tau^X) = \lim_{n \to \infty} Q(\tau_0^Z < \tau_n^X) = \lim_{n \to \infty} \mathbb{E}^P \left( X_{\tau_0^Z} \mathbb{1}_{\{\tau_0^Z < \tau_n^X\}} \right) = 0.
   \]
   Therefore,
   \[
   E(K, T) = Q \left( \rho_K \leq T < \tau_0^Z \right),
   \]
   \[
   A(K, T) = Q \left( \rho_K \leq T \wedge \tau^X \right) = Q \left( \rho_K \leq T \wedge \tau_0^Z \right) = Q \left( \rho_K \leq T \right),
   \]
   where the last equality follows from the fact that the last passage time of the level $1/K$ by $Z$ cannot be greater than its first hitting time of 0. We thus recover the formula for the European call option given in [22], Proposition 7, see also [32] for the case of the inverse
Bessel process of dimension greater than two. Especially, the American option premium is equal to

\[
A(K, T) - E(K, T) = Q(\rho K \leq T) - Q(\rho K \leq T < \tau_0^Z) = Q(\rho K \leq T, \tau_0^Z \leq T)
\]

which is just the default of the local \(P\)-martingale \(X\) or, in other words, the bubble of the stock \(X\) between 0 and \(T\).

(3) **\(X\) and \(Y\) are both strict local \(P\)-martingales: An example**

Let \(X\) and \(Y\) be the inverses of two independent BES(3)-processes under \(P\) and assume that \(X_0 = x \in \mathbb{R}_+\), while \(Y_0 = 1\). (Note that this normalization is different from the previous one. However, since the density of \(X\) resp. \(Y\) is explicitly known in this case, we can do calculations directly under \(P\). This allows us to point out some anomalies of the option value in the presence of strict local martingales).

We apply the formula for the European call option value written on an inverse BES(3)-process from Example 3.6 in [5] and integrate over \(Y\):

\[
E(K, T) = \int_0^\infty x \left[ \Phi \left( zK + x \sqrt{T} \right) - \Phi \left( zK - x \sqrt{T} \right) \right] P(Y_t \in dz)
\]

where

\[
P(Y_t \in dz) = 1 \frac{dz}{z^3 \sqrt{2\pi T}} \exp \left( -\frac{(1/z - 1)^2}{2T} \right) - \exp \left( -\frac{(1/z + 1)^2}{2T} \right).
\]

Since \(\mathbb{E}^P X_T \xrightarrow{T \to \infty} \frac{2}{\sqrt{2\pi T}}\), as shown in [15], the option value converges to a finite positive value as the initial stock price goes to infinity. Therefore, the convexity of the payoff function does not carry over to the option value. This anomaly for stock price bubbles has been noticed before by e.g. [5, 15]. We refer for the economic intuition of this phenomenon to [15], where a detailed analysis of stock and bond price bubbles modelled by the inverse BES(3)-process is done.

Furthermore, recall that by Jensen’s inequality the European exchange option value is increasing in maturity if \(X\) and \(Y\) are true martingales. However, in our example the option value is not increasing in maturity anymore: Since \(E(K, T) \leq \mathbb{E}^P X_T \xrightarrow{T \to \infty} 0\), the option value converges to zero as \(T \to \infty\). Taking \(Y \equiv 1\), this behaviour has been noticed before by e.g. [5, 15, 24, 28] and is also directly evident from the representation of \(E(K, T)\) in Theorem 5.1.

5.2. **Real-world pricing.** Here we want to give another interpretation of Theorem 5.1. Note that from a mathematical point of view we have only assumed that \(X\) and \(Y\) are strictly positive local \(P\)-martingales for the result. Above we have interpreted \(P\) as the risk-neutral probability and \(X, Y\) as two stock price processes. Now note that we have the identity \((X - KY)^+ = Y (\frac{1}{2} - K)^+\). This
motivates the following alternative financial setting: we take \( P \) to be the historical probability and assume that also \( P(Y_0 = 1) = 1 \). Normalizing the interest rate to be equal to zero, the process \( S := \frac{1}{Z} \) denotes the (discounted) stock price process, while \( Y \) is the density of a candidate for an equivalent local martingale measure (ELMM). Since \( Y \) and \( X = YS \) are both strictly positive local \( P \)-martingales, they are \( P \)-supermartingales and cannot reach infinity under \( P \). Thus, \( S = \frac{1}{Z} \) is also strictly positive under \( P \) and does not attain infinity under \( P \) either.

As before \( X \) and \( Y \) are both allowed to be either strict local or true \( P \)-martingales. While the question of whether \( X = YS \) is a true martingale or not is related to the existence of a stock price bubble as discussed earlier, the question of whether \( Y \) is a strict local martingale or not is connected to the absence of arbitrage. If \( Y \) is a true \( P \)-martingale, an ELMM for \( Z \) exists and the market satisfies NFLVR. However, as shown in [20] and explained in [1], even if \( Y \) is only a strict local martingale, a superhedging strategy for any contingent claim written on \( S \) exists. Therefore, the “normal” call option pricing formulas

\[
E(K,T) = \mathbb{E}^P (Y_T (S_T - K)^+) , \quad A(K,T) = \sup_{\sigma \in T_{0,T}} \mathbb{E}^P (Y_{\sigma} (S_{\sigma} - K)^+)
\]

are still reasonable when \( Y \) is only a strict local martingale. This pricing method is also known as "real-world pricing", since we cannot work under an ELMM directly, but must define the option value under the real-world measure, cf. [30]. Note that if \( Y \) is a true martingale, we can define an ELMM \( P^* \) for \( S \) via \( P^*|_{\mathcal{F}_t} = Y_t P|_{\mathcal{F}_t} \) and the market satisfies the NFLVR property. In this case we obtain the usual pricing formulas

\[
E(K,T) = \mathbb{E}^{P^*} (S_T - K)^+ \quad \text{resp.} \quad A(K,T) = \sup_{\sigma \in T_{0,T}} \mathbb{E}^{P^*} (S_{\sigma} - K)^+.
\]

Following [15] we can interpret the situation when \( Y \) is only a strict local martingale as the existence of a bond price bubble as opposed to the stock price bubble discussed above. This is motivated by the fact that the real-world price of a zero-coupon bond is strictly less than the (discounted) pay-off of one, if \( Y \) is a strict local martingale. Of course, it is possible to make a risk-free profit in this case via an admissible trading strategy. From Theorem 5.1 we have the following corollary:

**Corollary 5.4.** For all \( K, T \geq 0 \) the values of the European and American call option under real-world pricing are given by

\[
E(K,T) = \mathbb{E}^Q \left( \left( 1 - \frac{K}{S_{\tau X}} \right)^+ \mathbb{I}_{\{\rho_K^S \leq T < \tau X\}} \right), \quad A(K,T) = \mathbb{E}^Q \left( \left( 1 - \frac{K}{S_{\tau X}} \right)^+ \mathbb{I}_{\{\rho_K^S \leq \tau X \wedge T\}} \right),
\]

where \( \rho_K^S = \sup\{t \geq 0 \mid S_t = K\} \).

From the above formulas for the European and American call options it can easily be seen that their values are generally different, unless \( X = YS \) is a true \( P \)-martingale (in this case \( \tau X = \infty \).
Furthermore, note that we have the following formula for any bounded stopping time $T$:
\[
E(K,T) = E^P(X_T - K \cdot Y_T)^+ = E^Q(1 - KZ_T)^+ - E^Q\left(\mathbb{1}_{\{\tau^X_T \leq T\}}(1 - KZ_T)^+\right),
\]
where the second term equals $Q(\tau^X \leq T)$, if (A) holds. For $Y \equiv 1$ this decomposition of the European call value is shown in [28]. However, in general $Z$ is only a local martingale under $Q$. Therefore, the above formula is qualitatively different from the decomposition formulas in section ??.

Now we show that also the asymptotic behaviour of the European and American call option is unusual, when we allow $X$ and / or $Y$ to be strict local $P$-martingales. From the definition of the European call value we easily see that
\[
\lim_{K \to 0} E(K,T) = \lim_{K \to 0} A(K,T) = \lim_{K \to 0} Q(\rho_K^S \leq \tau^X \land T) = 1
\]
and
\[
\lim_{K \to \infty} A(K,T) = \lim_{K \to \infty} Q(\rho_K^S \leq \tau^X \land T, S_\tau^X = \infty) = \lim_{K \to \infty} Q(\rho_K^S \leq T, S_\tau^X = \infty)
\]
\[
= Q(\tau^X \leq T, Z_\tau^X = 0) = Q(Z_T = 0),
\]
which may be strictly positive and equals $Q(\tau^X \leq T)$ under (A).

For the asymptotics in $T$ we get
\[
\lim_{T \to \infty} E(K,T) = E^Q\left((1 - \frac{K}{S_\infty})^+ \mathbb{1}_{(\rho_K^S < \tau^X = \infty)}\right),
\]
\[
\lim_{T \to \infty} A(K,T) = E^Q\left((1 - \frac{K}{S_\tau^X})^+ \mathbb{1}_{(\rho_K^S < \tau^X)}\right),
\]
and from the definition of the call option it is also clear that
\[
\lim_{T \to 0} E(K,T) = \lim_{T \to 0} A(K,T) = (1 - K)^+.
\]

5.2.1. American option premium under real-world pricing. In this subsection we keep the notation and interpretation introduced in the last section. However, we do not assume that $Z$ and / or $X$ are continuous any more.

Lemma 5.5. Let $h : \mathbb{R}_+ \to \mathbb{R}_+$ be a Borel-measurable function such that $\lim_{x \to \infty} \frac{h(x)}{x} =: \eta$ exists in $\mathbb{R}_+$. Define $g : \mathbb{R}_+ \to \mathbb{R}_+$ via $g(x) = x \cdot h\left(\frac{1}{x}\right)$ for $x > 0$ and $g(0) = \eta$. We denote by
One may now naturally wonder whether, given two (or more) positive local $P$-martingales $X$ and $Y$ which are independent under $Q$, there exists a measure $Q$ which is only associated with the local $P$-martingale $Z$ and $Z_T = E(X_T)$ with payoff function $h$, and that $X$ and $Y$ are both strictly positive continuous local $P$-martingales with respect to the right-continuous augmentation $(\mathcal{F}_t)_{t \geq 0}$ of $(\tilde{\mathcal{F}}_t)_{t \geq 0}$ with $d\langle X \rangle_t = f_t dt, d\langle Y \rangle_t = g_t dt$ and $d\langle X,Y \rangle_t = h_t dt$. Suppose that $f_t g_t \neq h_t^2$ almost surely. Then there exists a measure $Q$ on $(\Omega, \mathcal{F}_{\tau \leq \tau_\varepsilon}, (\tilde{\mathcal{F}}_t)_{0 \leq t < \tau})$ under which $\frac{1}{X}$ and $\frac{1}{Y}$ are $Q$-martingales and $\frac{dP}{dQ} \bigg|_{\mathcal{F}_t \cap \mathcal{F}_{\tau_\varepsilon}} = \frac{1}{\varepsilon(M_t)}$, where

$$M_t = \int_0^t \frac{(f_s X_s - h_s Y_s)g_s}{Y_s X_s(f_sg_s - h_s^2)}dX_s + \int_0^t \frac{(g_s X_s - h_s Y_s)f_s}{Y_s X_s(f_sg_s - h_s^2)}dY_s.$$

**Proof.** For the European option value we have

$$E(h,T) = E^P(Y_T h(S_T)) = E^P \left( g(Z_T) \mathbb{1}_{\{\tau^X > T\}} \right) = E^Q \left( g(Z_T) \mathbb{1}_{\{\tau^X \leq T\}} \right).$$

And for the American option value we get

$$A(h,T) = \lim_{n \to \infty} E^P \left( Y_{T \wedge \tau_n^X} h(S_{T \wedge \tau_n^X}) \right) = \lim_{n \to \infty} E^Q \left( Z_{T \wedge \tau_n^X} h \left( \frac{1}{Z_{T \wedge \tau_n^X}} \right) \right) = \lim_{n \to \infty} E^Q \left( g(Z_{T \wedge \tau_n^X}) \right) = E^Q g(Z_T),$$

where the first equality is proven in [1] under the above stated assumptions on $h$ and the fourth equality follows by dominated convergence due to $g \leq 1$. □

Under the assumptions of Lemma 5.5 the American option premium is thus equal to

$$A(h,T) - E(h,T) = E^Q \left( g(Z_{T \wedge \tau_n^X}) \right).$$

For the European option value we have

$$E(h,T) = E^Q \left( g(Z_T) \mathbb{1}_{\{\tau^X \leq T\}} \right).$$

And for the American option value we get

$$A(h,T) = \lim_{n \to \infty} E^P \left( Y_{T \wedge \tau_n^X} h(S_{T \wedge \tau_n^X}) \right) = \lim_{n \to \infty} E^Q \left( Z_{T \wedge \tau_n^X} h \left( \frac{1}{Z_{T \wedge \tau_n^X}} \right) \right) = \lim_{n \to \infty} E^Q \left( g(Z_{T \wedge \tau_n^X}) \right) = E^Q g(Z_T),$$

where the first equality is proven in [1] under the above stated assumptions on $h$ and the fourth equality follows by dominated convergence due to $g \leq 1$. □

Note that Lemma 5.5 is a generalization of Theorem A1 in [3]. Indeed, if NFLVR is satisfied, then $Y$ is a true $P$-martingale and $Z_{\tau^X} = 0$ on $\{\tau^X < \infty\}$ by part 2(a) of Lemma 4.6. Thus,

$$A(h,T) = E(h,T) + g(0) \cdot Q \left( \tau^X \leq T \right) = E(h,T) + \gamma_X(0,T).$$

6. Multivariate Strictly Positive Local Martingales

So far the measure $Q$ defined in Theorem 1.8 above is only associated with the local $P$-martingale $X$ in the sense that $X_{\tau^X}.P|_{F_{\tau^X}} = Q_{|_{F_{\tau^X}}}$ for all $n \in \mathbb{N}$ and that $\frac{1}{X}$ is a true martingale under $Q$. One may now naturally wonder whether, given two (or more) positive local $P$-martingales $X$ and $Y$, there exists a measure $Q$, under which $\frac{1}{X}$ and $\frac{1}{Y}$ are both true martingales. Obviously, this is the case, if $X$ and $Y$ are independent under $P$. Moreover, if $X$ and $Y$ are continuous, this question has a positive answer as is shown in the following theorem:

**Theorem 6.1.** Assume that the filtration of $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, P)$ is a standard system. Let $X, Y$ be both strictly positive continuous local $P$-martingales with respect to the right-continuous augmentation $(\mathcal{F}_t)_{t \geq 0}$ of $(\tilde{\mathcal{F}}_t)_{t \geq 0}$ with $d\langle X \rangle_t = f_t dt, d\langle Y \rangle_t = g_t dt$ and $d\langle X, Y \rangle_t = h_t dt$. Suppose that $f_t g_t \neq h_t^2$ almost surely. Then there exists a measure $Q$ on $(\Omega, \mathcal{F}_{\tau \leq \tau_\varepsilon}, (\tilde{\mathcal{F}}_t)_{0 \leq t < \tau})$, under which $\frac{1}{X}$ and $\frac{1}{Y}$ are $Q$-martingales and $\frac{dP}{dQ} \bigg|_{\mathcal{F}_t \cap \mathcal{F}_{\tau_\varepsilon}} = \frac{1}{\varepsilon(M_t)}$, where

$$M_t = \int_0^t \frac{(f_s X_s - h_s Y_s)g_s}{Y_s X_s(f_sg_s - h_s^2)}dX_s + \int_0^t \frac{(g_s X_s - h_s Y_s)f_s}{Y_s X_s(f_sg_s - h_s^2)}dY_s.$$
and $\tau^\xi$ is the explosion time of $\mathcal{E}(M)$.

Proof. The stochastic exponential $\mathcal{E}(M)$ is a continuous local $\mathbb{P}$-martingale with localizing sequence

$$\tau^\xi_n := \inf\{t \geq 0 : \mathcal{E}(M)_t > n\} \wedge n.$$ 

We define a consistent family of probability measures $Q_n$ on $\mathcal{F}_{\tau^\xi_n}$ by

$$\frac{dQ_n}{d\mathbb{P}}\bigg|_{\mathcal{F}_{\tau^\xi_n}} = \mathcal{E}(M)_{\tau^\xi_n}, \quad n \in \mathbb{N}.$$ 

Using the same trick as in the proof of Theorem 1.8, we restrict each measure $Q_n$ to $\mathcal{F}_{\tau^\xi_n}$. Since $(\mathcal{F}_{\tau^\xi_n})_{n \in \mathbb{N}}$ is a standard system by Lemma 1.7, there exists a unique measure $Q$ on $\mathcal{F}_{\tau^\xi_n}$, such that $Q|_{\mathcal{F}_{\tau^\xi_n}} = Q_n$ for all $n \in \mathbb{N}$. For any stopping time $S$ and $A \in \mathcal{F}_S$ we get

$$Q(S < \tau^\xi_n, A) = \mathbb{E}^Q(\mathcal{E}(M)_{S \wedge \tau^\xi_n} \mathbb{1}_{\{S < \tau^\xi_n, A\}}).$$

Taking $n \to \infty$ results in

$$Q(S < \tau^\xi, A) = \mathbb{E}^Q(\mathcal{E}(M)_S \mathbb{1}_{\{S < \infty, A\}}).$$

It follows that $\mathbb{P}$ is locally absolutely continuous with respect to $Q$ before $\tau^\xi$. Next, according to Girsanov’s theorem applied on $\mathcal{F}_{\tau^\xi_n}$

$$N_{t \wedge \tau^\xi_n} := X^{\tau^\xi}_t - \langle M^{\tau^\xi}_n, X^{\tau^\xi}_n \rangle_t = X^{\tau^\xi}_t - \int_0^{t \wedge \tau^\xi_n} \left( f_s Y_s - h_s X_s \right) g_s d\langle X \rangle_s - \int_0^{t \wedge \tau^\xi_n} \frac{g_s X_s - h_s Y_s}{Y_s X_s (f_s g_s - h_s^2)} ds \left( X_s Y_s - f_s g_s Y_s - h_s X_s \right) d\langle X, Y \rangle_s$$

is a local $Q$-martingale. We apply Itô’s formula:

$$\frac{1}{X_{t \wedge \tau^\xi_n}} = \frac{1}{X_0} - \int_0^{t \wedge \tau^\xi_n} \frac{dX_s}{X_s^2} + \int_0^{t \wedge \tau^\xi_n} \frac{d\langle X \rangle_s}{X_s^3} = \frac{1}{X_0} - \int_0^{t \wedge \tau^\xi_n} \frac{dN_s}{X_s^2} + \frac{1}{X_0} - \int_0^{t \wedge \tau^\xi_n} \int_0^{m \wedge \tau^\xi_n} \frac{f_s}{X_s^3} ds d\langle X \rangle_s.$$

Thus, $\frac{1}{X_{t \wedge \tau^\xi_n}}$ is a local $Q$-martingale for all $n \in \mathbb{N}$. Since $\frac{1}{X}$ is continuous, $(\tau^{1/X}_m)_{m \in \mathbb{N}}$ is a localizing sequence for $\frac{1}{X_{t \wedge \tau^\xi_n}}$ on $(\Omega, \mathcal{F}_{\tau^\xi_n}, Q)$ for all $n \in \mathbb{N}$, where

$$\tau^{1/X}_m := \inf\{t \geq 0 : \frac{1}{X_t} > m\} \wedge m, \quad \tau^{1/X}_m := \lim_{m \to \infty} \tau^{1/X}_m.$$

Moreover, we have

$$Q(\tau^{1/X} < \tau^\xi) = \lim_{n \to \infty} Q(\tau^{1/X} < \tau^\xi_n) = \lim_{n \to \infty} \mathbb{E}^Q(\mathcal{E}(M)_{\tau^\xi_n} \mathbb{1}_{\{\tau^{1/X} < \tau^\xi_n\}}) = 0,$$

which yields that

$$\frac{1}{X_0} = \lim_{m \to \infty} \mathbb{E}^Q\left( \frac{1}{X_{t \wedge \tau^\xi_n \wedge \tau^{1/X}_m}} \right) = \lim_{m \to \infty} \mathbb{E}^Q\left( \frac{1}{X_{t \wedge \tau^\xi_n}} \mathbb{1}_{\{\tau^{1/X}_m < \tau^\xi_n\}} \right) + \lim_{m \to \infty} \mathbb{E}^Q\left( \frac{1}{X_{t \wedge \tau^\xi_n}} \mathbb{1}_{\{\tau^{1/X}_m \geq \tau^\xi_n\}} \right)$$

$$= \lim_{m \to \infty} \mathbb{E}^Q\left( \frac{1}{X_{t \wedge \tau^\xi_n}} \right) + \lim_{m \to \infty} \mathbb{E}^Q\left( \frac{1}{X_{t \wedge \tau^\xi_n}} \mathbb{1}_{\{\tau^\xi_n \geq \tau^{1/X}_m\}} \right).$$
Since $\frac{1}{X}$ is a $P$-submartingale,
\[
\lim_{m \to \infty} \mathbb{E}^Q\left( \frac{1}{X_{\tau_n \wedge \tau_{\frac{m}{n}}}} 1_{\{\tau_{\frac{m}{n}} \geq \tau_n^{1/2}\}} \right) = \lim_{m \to \infty} \mathbb{E}^P\left( \frac{\mathcal{E}(M)_{\tau_{\frac{m}{n}}} \mathbb{E}_2^P \{\tau_{\frac{m}{n}} \geq \tau_n^{1/2}\}}{X_{\tau_n \wedge \tau_{\frac{m}{n}}}} \right) \leq \lim_{m \to \infty} \mathbb{E}^P\left( \frac{n}{X_{\tau_n \wedge \tau_{\frac{m}{n}}}} 1_{\{\tau_{\frac{m}{n}} \geq \tau_n^{1/2}\}} \right) = 0.
\]
Therefore, $\frac{1}{X_{\tau_n}}$ is actually a $Q$-martingale, which shows that $\frac{1}{X}$ is a $Q$-martingale on the stochastic interval $[0, \tau_{\frac{m}{n}})$.

For $\frac{1}{X}$ the claim follows by analogous calculations. □

**Remark 6.2.** The above theorem deals with two strictly positive local $P$-martingales. It is however obvious that one can get a similar result for $n \geq 2$ strictly positive local martingales.

We briefly want to describe a different approach focusing on “conformal local martingales” in $\mathbb{R}^d$, $d > 2$, which is dealt with in [28]. However, in [28] the authors make the restriction that the conformal local martingale does not enter some compact neighborhood of the origin in $\mathbb{R}^d$. Using simple localization arguments as in Theorem 1.8 above, one can get rid off this assumption which seems somehow inappropriate when dealing with stock price processes. This yields the following extended version of Lemma 12 in [28]. We denote by $| \cdot |$ the Euclidean norm in $\mathbb{R}^d$.

**Definition 6.3.** A continuous local martingale $X$, taking values in $\mathbb{R}^d$, is called a conformal local martingale on $(\Omega, (\mathcal{F}_t)_{t \geq 0}, P)$, if $(X^i, X^j) = (X^1)_{i=j}$ $P$-almost surely for all $1 \leq i, j \leq d$.

**Theorem 6.4.** Assume that the filtration of $(\Omega, (\mathcal{F}_t)_{t \geq 0}, P)$ is a standard system and denote by $(\mathcal{F}_t)_{t \geq 0}$ its right-continuous augmentation. Let $X = (X^1, \ldots, X^d)$ be a conformal local $P, (\mathcal{F}_t)$-martingale, where each $X^i$ is non-negative and $X \neq 0$ $P$-almost surely. Suppose $X_0 = x_0$ with $|x_0| = 1$. Define $\tau := \inf\{t \geq 0 \mid |X_t| = 0\}$. Then there exists a measure $Q$ on $\mathcal{F}_\tau$, such that $Y := \frac{X}{|X|}$ is a non-negative $Q$-martingale with $(Y^i, Y^j) = (Y^1)_{i=j}$ $Q$-a.s. for all $1 \leq i, j \leq d$.

**Proof.** Define $\tau_n := \inf\{t \geq 0 \mid |X_t| \leq \frac{1}{n}\}$. As in Lemma 11 in [28] it follows that $(|X_{\tau_n \wedge \tau}|^2)^{1/2}_{t \geq 0}$ is a uniformly integrable $P$-martingale for all $n \in \mathbb{N}$, because $| \cdot |^{2-d}$ is harmonic. We define a consistent family of probability measures $Q_n$ on $\mathcal{F}_{\tau_n}$ by
\[
\frac{dQ_n}{dP} \bigg|_{\mathcal{F}_{\tau_n}} = |X_{\tau_n}|^{2-d}, \quad n \in \mathbb{N}.
\]

Using the same trick as in the proof of Theorem 1.8 we restrict each measure $Q_n$ to $\mathcal{F}_{\tau_n}$. Since $(\mathcal{F}_{\tau_n})_{n \in \mathbb{N}}$ is a standard system, there exists a unique measure $Q$ on $\mathcal{F}_\tau$, such that $Q|_{\mathcal{F}_{\tau_n}} = Q_n$ for all $n \in \mathbb{N}$. For any stopping time $S$ we thus get
\[
Q(S < \tau_n) = \mathbb{E}^P\left(|X_{\tau_n}|^{2-d} 1_{\{S < \tau_n\}}\right) = \mathbb{E}^P\left(|X_S|^{2-d} 1_{\{S < \tau_n\}}\right).
\]
Choosing $S = t < \infty$ and taking $n \to \infty$ results in
\[
Q(t < \tau) = \mathbb{E}^P|X_t|^{2-d}.
\]
Therefore, $P$ is locally absolutely continuous to $Q$ before $\tau$.

From Lemma 12 in [28] we know that $\frac{X_{\tau_n \wedge \tau}}{|X_{\tau_n \wedge \tau}|^2}$ is a conformal $Q_n$-martingale and thus $Y$ is a $Q$-martingale up to time $\tau$. □
Appendix A. Condition (P)

In Theorem 1.1 we mentioned the condition (P), which was introduced in Definition 4.1 in [27] following [29] as follows:

**Definition A.1.** Let \((\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0})\) be a filtered measurable space, such that \(\mathcal{F}\) is the \(\sigma\)-algebra generated by \((\mathcal{F}_t)_{t \geq 0}\): \(\mathcal{F} = \bigvee_{t \geq 0} \mathcal{F}_t\). We shall say that the property (P) holds if and only if \((\mathcal{F}_t)_{t \geq 0}\) enjoys the following conditions:

- For all \(t \geq 0\), \(\mathcal{F}_t\) is generated by a countable number of sets;
- For all \(t \geq 0\), there exists a Polish space \(\Omega_t\), and a surjective map \(\pi_t\) from \(\Omega\) to \(\Omega_t\), such that \(\mathcal{F}_t\) is the \(\sigma\)-algebra of the inverse images, by \(\pi_t\), of Borel sets in \(\Omega_t\), and such that for all \(B \in \mathcal{F}_t\), \(\omega \in \Omega\), \(\pi_t(\omega) \in \pi_t(B)\) implies \(\omega \in B\);
- If \((\omega_n)_{n \geq 0}\) is a sequence of elements of \(\Omega\), such that for all \(N \geq 0\),
  \[
  \bigcap_{n \geq 0} A_n(\omega_n) \neq \emptyset,
  \]
  where \(A_n(\omega_n)\) is the intersection of the sets in \(\mathcal{F}_n\) containing \(\omega_n\), then:
  \[
  \bigcap_{n \geq 0} A_n(\omega_n) \neq \emptyset.
  \]

Appendix B. Doob h-transform

Doob h-transforms are defined in the theory of Markov processes as follows:

Let \(X\) be a Markov process starting from \(x\) with values in \(I\) and a cemetery state \(\delta\) under the measure \(P_x\). Let \(\mathcal{F}_t = \sigma(X_s; s \leq t)\) for all \(t \geq 0\), and let \(h\) be an excessive function for \(X\), i.e.
\[
E_x h(X_t) \leq h(x) \quad \forall \, x \in I, \quad \forall \, t \geq 0, \quad E_x h(X_t) \to h(x) \quad \forall \, x \in I \text{ as } t \to 0.
\]
Then a new sub-Markovian measure can be defined via
\[
P^h_x|_{\mathcal{F}_t} = \frac{h(X_t)}{h(x)} \cdot P_x|_{\mathcal{F}_t}, \quad t < \zeta := \inf\{t \geq 0\} \ | \ X_t = \delta\}.
\]
Then, \(X\) considered as a process under \(P^h\) is called the Doob h-transform of \(X\).

If one takes \(h\) to be the identity and requires \(X\) to be a Markovian strictly positive local martingale, we have
\[
\mathbb{1}_{\{t < \zeta\}} P^h_x|_{\mathcal{F}_t} = \frac{X_t}{x} \cdot P_x|_{\mathcal{F}_t},
\]
\(i.e.\ \text{under } P^h_x \text{ the process } X \text{ is killed at } \zeta, \text{ which happens with positive probability, if } X \text{ is a strict local martingale.}
Note however that we do not assume that the process is Markov in the analysis done in this paper.
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